

Appendices

Appendix 1: Derivations for the Time-varying Inefficiency Model

Consider the frontier production function¹

$$Y_{it} = x_{it}\beta + E_{it} \quad (A1.1)$$

where

$$E_{it} = V_{it} - \eta_{it}U_i \quad (A1.2)$$

and

$$\eta_{it} = e^{-\eta(t-T)}, \quad t \in \mathcal{T}(i); \quad i = 1, 2, \dots, N. \quad (A1.3)$$

It is assumed that the V_{it} s are independent and identically distributed $N(0, \sigma_v^2)$ random variables, independent of the U_i s, which are assumed to be non-negative truncations of the $N(\mu, \sigma^2)$ distribution.

The density function for U_i is

$$f_{U_i}(u_i) = \frac{\exp\left[-\frac{1}{2}(u_i - \mu)^2 / \sigma^2\right]}{(2\pi)^{1/2} \sigma [1 - \Phi(-\mu / \sigma)]}, \quad u_i \geq 0, \quad (A1.4)$$

where $\Phi(\cdot)$ represents the distribution function for the standard normal random variable.

It can be shown that the mean and variance of U_i are²

¹ In the frontier model (3.2), the notation, Y_{it} , represented the actual production at the time of the t -th observation for the i -th firm. However, given that (3.2) involves a Cobb-Douglas or transcendental logarithmic model, then Y_{it} and x_{it} in this Appendix would represent logarithms of output and input values, respectively.

² We prefer not to use the notation, $\sigma_{U_i}^2$, for the variance of the normal distribution which is truncated (at zero) to obtain the distribution of the non-negative technical inefficiency effects, because this variance is not the variance of U_i . For the case of the half-normal distribution the variance of U_i is $\sigma^2(\pi-2)/\pi$. This fact needs to be kept in mind in the interpretation of empirical results for the stochastic frontier model.

$$E(U_i) = \mu + \sigma \left\{ \phi(-\mu / \sigma) / [1 - \Phi(-\mu / \sigma)] \right\} \quad (A1.5)$$

and

$$\text{Var}(U_i) = \sigma^2 \left\{ 1 - \frac{\phi(-\mu / \sigma)}{1 - \Phi(-\mu / \sigma)} \left[\frac{\mu}{\sigma} + \frac{\phi(-\mu / \sigma)}{1 - \Phi(-\mu / \sigma)} \right] \right\}, \quad (A1.6)$$

where $\phi(\cdot)$ represents the density function for the standard normal distribution.

From the joint density function for U_i and V_i , where V_i represents the $(T_i \times 1)$ vector of the V_{it} s associated with the T_i observations for the i -th firm, it follows readily that the joint density function for U_i and E_i , where E_i is the $(T_i \times 1)$ vector of the values of $E_{it} \equiv V_{it} - \eta_{it}U_i$, is

$$f_{U_i, E_i}(u_i, e_i) = \frac{\exp - \frac{1}{2} \left\{ \left[(u_i - \mu)^2 / \sigma^2 \right] + \left[(e_i + \eta_i u_i)' (e_i + \eta_i u_i) / \sigma_v^2 \right] \right\}}{(2\pi)^{(T_i+1)/2} \sigma \sigma_v^{T_i} [1 - \Phi(-\mu / \sigma)]} \quad (A1.7)$$

where e_i is a possible value for the random vector, E_i .

The density function for E_i , obtained by integrating $f_{U_i, E_i}(u_i, e_i)$ with respect to the range for U_i , namely $u_i \geq 0$, is

$$f_{E_i}(e_i) = \frac{[1 - \Phi(-\mu_i^* / \sigma_i^*)] \exp - \frac{1}{2} \left\{ (e_i' e_i / \sigma_v^2) + (\mu / \sigma)^2 - (\mu_i^* / \sigma_i^*)^2 \right\}}{(2\pi)^{T_i/2} \sigma_v^{(T_i-1)} [\sigma_v^2 + \eta_i' \eta_i \sigma^2]^{1/2} [1 - \Phi(-\mu / \sigma)]} \quad (A1.8)$$

$$\text{where} \quad \mu_i^* \equiv \frac{\mu \sigma_v^2 - \eta_i' e_i \sigma^2}{\sigma_v^2 + \eta_i' \eta_i \sigma^2} \quad (A1.9)$$

$$\text{and} \quad \sigma_i^{*2} \equiv \frac{\sigma_v^2 \sigma^2}{\sigma_v^2 + \eta_i' \eta_i \sigma^2}. \quad (A1.10)$$

From the above results, it follows that the conditional density function of U_i , given that the random vector, E_i , has value, e_i , is

$$f_{U_i|E_i=e_i}(u_i) = \frac{\exp - \frac{1}{2} \left[(\mu_i - \mu_i^*) / \sigma_i^* \right]^2}{(2\pi)^{1/2} \sigma_i^* [1 - \Phi(-\mu_i^* / \sigma_i^*)]}, \quad u_i \geq 0. \quad (A1.11)$$

This is the density function of the positive truncation of the $N(\mu_i^* / \sigma_i^{*2})$ distribution.

Since the conditional expectation of $\exp(-\eta_{it}U_i)$, given $E_i = e_i$, is defined by

$$E\{\exp(-\eta_{it}U_i|E_i = e_i)\} = \int_0^\infty \exp(-\eta_{it}u_i) f_{U_i|E_i=e_i}(u_i) du_i ,$$

the result of equation (3.3) is obtained by straightforward integral calculus.

If the frontier production function (A1.1)-(A1.3) is appropriate for production, expressed in the original units of output, then the prediction of the technical efficiency of the i -th firm at the time of the t -th observation, $TE_{it} = 1 - (\eta_{it}U_i/x_{it}\beta)$, requires the conditional expectation of U_i , given $E_i = e_i$. This can be shown to be

$$E(U_i|E_i = e_i) = \mu_i^* + \sigma_i^* \left\{ \phi(-\mu_i^* / \sigma_i^*) / \left[1 - \Phi(-\mu_i^* / \sigma_i^*) \right] \right\} \quad (A1.12)$$

where μ_i^* and σ_i^{*2} are defined by equations (A1.9) and (A1.10), respectively.

The density function for Y_i , the $(T_i \times 1)$ random vector of Y_{it} s for the i -th firm, is obtained from (A1.8) by substituting $(y_i - x_i\beta)$ for e_i , where x_i is the $(T_i \times k)$ matrix of x_{it} s for the i -th firm, where k is the dimension of the vector, β . The logarithm of the likelihood function for the sample observations, $y \equiv (y'_1, y'_2, \dots, y'_N)'$, is thus

$$\begin{aligned} L^*(\theta^*; y) = & -\frac{1}{2} \left(\sum_{i=1}^N T_i \right) \ln(2\pi) - \frac{1}{2} \sum_{i=1}^N (T_i - 1) \ln(\sigma_v^2) - \frac{1}{2} \sum_{i=1}^N \ln(\sigma_v^2 + \eta_i' \eta_i \sigma^2) \\ & - N \ln[1 - \Phi(-\mu / \sigma)] + \sum_{i=1}^N \ln[1 - \Phi(-\mu_i^* / \sigma_i^*)] \\ & - \frac{1}{2} \sum_{i=1}^N \left[(y_i - x_i\beta)' (y_i - x_i\beta) / \sigma_v^2 \right] - \frac{1}{2} N(\mu / \sigma)^2 + \frac{1}{2} \sum_{i=1}^N (\mu_i^* / \sigma_i^*)^2 \end{aligned} \quad (A1.13)$$

where $\theta^* \equiv (\beta', \sigma_v^2, \sigma^2, \mu, \eta)'$.

Using the reparameterisation of the model, suggested by Battese and Corra (1977), where $\sigma_v^2 + \sigma^2 = \sigma_s^2$ and $\gamma = \sigma^2 / \sigma_s^2$, the logarithm of the likelihood function is expressed by

$$\begin{aligned}
L^*(\theta; y) = & -\frac{1}{2} \left(\sum_{i=1}^N T_i \right) \left\{ \ell n(2\pi) + \ell n(\sigma_s^2) \right\} - \frac{1}{2} \sum_{i=1}^N (T_i - 1) \ell n(1 - \gamma) \\
& - \frac{1}{2} \sum_{i=1}^N \ell n[1 + (\eta'_i \eta_i - 1)\gamma] - N \ell n[1 - \Phi(-z)] - \frac{1}{2} N z^2 \\
& + \sum_{i=1}^N \ell n[1 - \Phi(-z_i^*)] + \frac{1}{2} \sum_{i=1}^N z_i^{*2} - \frac{1}{2} \sum_{i=1}^N (y_i - x_i \beta)' (y_i - x_i \beta) / (1 - \gamma) \sigma_s^2,
\end{aligned} \tag{A1.14}$$

where $\theta \equiv (\beta', \sigma_s^2, \gamma, \mu, \eta)'$, $z \equiv \mu/(\gamma \sigma_s^2)^{1/2}$ and

$$z_i^* = \frac{\mu(1 - \gamma) - \gamma \eta'_i (y_i - x_i \beta)}{\left\{ \gamma(1 - \gamma) \sigma_s^2 [1 + (\eta'_i \eta_i - 1)\gamma] \right\}^{1/2}}.$$

The partial derivations of the loglikelihood function (A1.14) with respect to the parameters, β , σ_s^2 , γ , μ and η , are given by

$$\begin{aligned}
\frac{\partial L^*}{\partial \beta} &= \sum_{i=1}^N x_i' (y_i - x_i \beta) \left[(1 - \gamma) \sigma_s^2 \right]^{-1} \\
&+ \sum_{i=1}^N \left[\frac{\phi(-z_i^*)}{1 - \Phi(-z_i^*)} + z_i^* \right] \gamma x_i' \eta_i \left\{ \gamma(1 - \gamma) \sigma_s^2 [1 + (\eta'_i \eta_i - 1)\gamma] \right\}^{-1/2} \\
\frac{\partial L^*}{\partial \sigma_s^2} &= -\frac{1}{2\sigma_s^2} \left\{ \sum_{i=1}^N T_i - N \left[\frac{\phi(-z)}{1 - \Phi(-z)} + z \right] z + \sum_{i=1}^N \left[\frac{\phi(-z_i^*)}{1 - \Phi(-z_i^*)} + z_i^* \right] z_i^* \right. \\
&\quad \left. - \sum_{i=1}^N (y_i - x_i \beta)' (y_i - x_i \beta) \left[(1 - \gamma) \sigma_s^2 \right]^{-1} \right\} \\
\frac{\partial L^*}{\partial \gamma} &= \frac{(1 - \gamma)^{-1}}{2} \sum_{i=1}^N (T_i - 1) - \frac{1}{2} \sum_{i=1}^N (\eta'_i \eta_i - 1) [1 + (\eta'_i \eta_i - 1)\gamma]^{-1} \\
&+ \frac{N}{2} \left[\frac{\phi(-z)}{1 - \Phi(-z)} + z \right] z \gamma^{-1} + \sum_{i=1}^N \left[\frac{\phi(-z_i^*)}{1 - \Phi(-z_i^*)} + z_i^* \right] \frac{\partial z_i^*}{\partial \gamma} \\
&- \frac{1}{2} \sum_{i=1}^N (y_i - x_i \beta)' (y_i - x_i \beta) \left[(1 - \gamma) \sigma_s^2 \right]^{-2}
\end{aligned}$$

$$\frac{\partial L^*}{\partial \mu} = -\frac{N}{(\gamma \sigma_s^2)^{1/2}} \left[\frac{\phi(-z)}{1 - \Phi(-z)} + z \right] + \sum_{i=1}^N \left[\frac{\phi(-z_i^*)}{1 - \Phi(-z_i^*)} + z_i^* \right] \\ \times \frac{(1 - \gamma)}{\{\gamma(1 - \gamma) \sigma_s^2 [1 + (\eta_i' \eta_i - 1) \gamma]\}^{1/2}}$$

$$\frac{\partial L^*}{\partial \eta} = \sum_{i=1}^N \left[\frac{\phi(-z_i^*)}{1 - \Phi(-z_i^*)} + z_i^* \right] \frac{\partial z_i^*}{\partial \eta} - \frac{\gamma}{2} \sum_{i=1}^N \frac{\partial \eta_i' \eta_i}{\partial \eta} [1 + (\eta_i' \eta_i - 1) \gamma]^{-1}$$

$$\frac{\partial z_i^*}{\partial \gamma} = -\frac{[\mu + \eta_i'(y_i - x_i \beta)]}{\sigma_s \{\gamma(1 - \gamma) [1 + (\eta_i' \eta_i - 1) \gamma]\}^{1/2}} \\ - \frac{1}{2} \frac{[\mu(1 - \gamma) - \eta_i'(y_i - x_i \beta)] [(1 - 2\gamma) + (\eta_i' \eta_i - 1) \gamma(2 - 3\gamma)]}{\sigma_s \{\gamma(1 - \gamma) [1 + (\eta_i' \eta_i - 1) \gamma]\}^{3/2}}$$

$$\frac{\partial z_i^*}{\partial \eta} = \frac{\gamma \sum_{t \in \partial(i)} (t - T) e^{-\eta(t-T)} (y_{it} - x_{it} \beta)}{\{\gamma(1 - \gamma) \sigma_s^2 [1 + (\eta_i' \eta_i - 1) \gamma]\}^{1/2}} - \frac{[\mu(1 - \gamma) - \eta_i'(y_i - x_i \beta)] \frac{1}{2} \gamma^2 (1 - \gamma) \sigma_s^2 \frac{\partial \eta_i' \eta_i}{\partial \eta}}{\{\gamma(1 - \gamma) \sigma_s^2 [1 + (\eta_i' \eta_i - 1) \gamma]\}^{3/2}}$$

$$\text{and} \quad \frac{\partial \eta_i' \eta_i}{\partial \eta} = -2 \sum_{t \in \partial(i)} (t - T) e^{-2\eta(t-T)} \quad \text{if } \eta \neq 0.$$

Appendix 2: Derivations for the Stochastic Frontier and Inefficiency Model

For simplicity of presentation of results in this Appendix, we assume that the stochastic frontier and inefficiency model is expressed by

$$Y_{it} = x_{it}\beta + E_{it} \quad (A2.1)$$

and

$$E_{it} = V_{it} - U_{it} , \quad (A2.2)$$

where $i = 1, 2, \dots, N$, and $t = 1, 2, \dots, T$.

Thus, given the frontier production function of equation (4.1), Y_{it} in equation (A2.1) is, in fact, the logarithm of the production for the i -th firm in the t -th time period. Apart from introducing the notation, E_{it} , for the difference between V_{it} and U_{it} , all other variables and parameters in equations (A2.1) and (A2.2) are the same as in equations (4.1) and (4.2).

The density functions for V_{it} and U_{it} are

$$f_V(v) = \frac{\exp(-\frac{1}{2}v^2 / \sigma_V^2)}{\sqrt{2\pi} \sigma_V}, \quad -\infty < v < \infty \quad (A2.3)$$

and

$$f_U(u) = \frac{\exp\left[-\frac{1}{2}(u - z\delta)^2 / \sigma^2\right]}{\sqrt{2\pi} \sigma \Phi[z\delta / \sigma]}, \quad u \geq 0, \quad (A2.4)$$

where the subscripts, i and t , are omitted for convenience in the presentation.

The joint density function for $E = V - U$ and U is

$$\begin{aligned} f_{E,U}(e, u) &= \frac{\exp\left\{-\frac{1}{2}\left[\left(\frac{e+u}{\sigma_V}\right)^2 + \left(\frac{u-z\delta}{\sigma}\right)^2\right]\right\}}{2\pi \sigma \sigma_V \Phi[z\delta / \sigma]}, \quad u \geq 0 \\ &= \frac{\exp\left\{-\frac{1}{2}\left[\left(\frac{u-\mu_*}{\sigma_*}\right)^2 + \left(\frac{e^2}{\sigma_V^2}\right) + \left(\frac{z\delta}{\sigma}\right)^2 - \left(\frac{\mu_*}{\sigma_*}\right)^2\right]\right\}}{2\pi \sigma \sigma_V \Phi[z\delta / \sigma]} \end{aligned} \quad (A2.5a)$$

or, alternatively,

$$f_{E,U}(e,u) = \frac{\exp\{-\frac{1}{2}\left[\left(u - \mu_*$$

where

$$\mu_* = \frac{\sigma_v^2 z\delta - \sigma^2 e}{\sigma_v^2 + \sigma^2} \quad (A2.6)$$

and

$$\sigma_*^2 = \sigma^2 \sigma_v^2 / (\sigma^2 + \sigma_v^2) . \quad (A2.7)$$

Thus the density function for $E = V - U$ is

$$\begin{aligned} f_E(e) &= \frac{\exp\{-\frac{1}{2}\left[(e^2 / \sigma_v^2) + (z\delta / \sigma)^2 - (\mu_* / \sigma_*)^2\right]\}}{\sqrt{2\pi} \sigma_v \sigma \Phi[z\delta / \sigma]} \int_0^\infty \frac{\exp\{-\frac{1}{2}\left[(u - \mu_*)^2 / \sigma_*^2\right]\}}{\sqrt{2\pi}} du \\ &= \frac{\exp\{-\frac{1}{2}\left[(e^2 / \sigma_v^2) + (z\delta / \sigma)^2 - (\mu_* / \sigma_*)^2\right]\}}{\sqrt{2\pi} (\sigma^2 + \sigma_v^2)^{1/2} [\Phi(z\delta / \sigma) / \Phi(\mu_* / \sigma_*)]} , \end{aligned} \quad (A2.8a)$$

or, alternatively,

$$f_E(e) = \frac{\exp\left\{-\frac{1}{2}(e + z\delta)^2 / (\sigma_v^2 + \sigma^2)\right\}}{\sqrt{2\pi} (\sigma_v^2 + \sigma^2)^{1/2} [\Phi(z\delta / \sigma) / \Phi(\mu_* / \sigma_*)]} . \quad (A2.8b)$$

The conditional density function for U given $E = e$ is thus

$$f_{U|E=e}(u) = \frac{\exp\{-\frac{1}{2}\left[(u - \mu_*)^2 / \sigma_*^2\right]\}}{\sqrt{2\pi} \sigma_* \Phi(\mu_* / \sigma_*)} , \quad u \geq 0 . \quad (A2.9)$$

It can be shown that the conditional expectation of e^{-U} , given $E = e$, is

$$E(e^{-U} | E = e) = \left\{ \exp\left[-\mu_* + \frac{1}{2} \sigma_*^2\right] \right\} \left\{ \Phi\left[(\mu_* / \sigma_*) - \sigma^*\right] / \Phi(\mu_* / \sigma_*) \right\} . \quad (A2.10)$$

The density function for the production value, Y_{it} , in equation (A2.1), is most conveniently given using the expression in equation (A2.8b),

$$f_{Y_{it}}(y_{it}) = \frac{\exp\left\{-\frac{1}{2} \frac{(y_{it} - x_{it}\beta + z_{it}\delta)^2}{\sigma_v^2 + \sigma^2}\right\}}{\sqrt{2\pi} (\sigma_v^2 + \sigma^2)^{1/2} [\Phi(d_{it}) / \Phi(d_{it}^*)]} \quad (A2.11)$$

where $d_{it} = z_{it}\delta / \sigma$, $d_{it}^* = \mu_{it}^* / \sigma_*$ and $\mu_{it}^* = [\sigma_v^2 z_{it}\delta - \sigma^2(y_{it} - x_{it}\beta)] / (\sigma_v^2 + \sigma^2)$.

Given that there are T_i observations obtained for the i -th firm, where $1 \leq T_i \leq T$, and $Y_i \equiv (Y_{i1}, Y_{i2}, \dots, Y_{iT_i})'$ denotes the vector of the T_i production values in equation (A2.1), then the logarithm of the likelihood function for the sample observations, $y \equiv (y_1', y_2', \dots, y_T')'$, is

$$\begin{aligned} L^*(\theta^*; y) = & -\frac{1}{2} \left(\sum_{i=1}^N T_i \right) \left\{ \ln 2\pi + \ln(\sigma^2 + \sigma_v^2) \right\} \\ & - \frac{1}{2} \sum_{i=1}^N \sum_{t=1}^{T_i} \left[(y_{it} - x_{it}\beta + z_{it}\delta)^2 / (\sigma_v^2 + \sigma^2) \right] \\ & - \sum_{i=1}^N \sum_{t=1}^{T_i} \left[\ln \Phi(d_{it}) - \ln \Phi(d_{it}^*) \right] \end{aligned} \quad (A2.12)$$

where $\theta^* = (\beta', \delta', \sigma_v^2, \sigma^2)'$.

Using the re-parameterisation of the model involving the parameters, $\sigma_s^2 \equiv \sigma_v^2 + \sigma^2$ and $\gamma \equiv \sigma^2 / \sigma_s^2$, the logarithm of the likelihood function is expressed by

$$\begin{aligned} L^*(\theta; y) = & -\frac{1}{2} \left(\sum_{i=1}^N T_i \right) \left\{ \ln 2\pi + \ln \sigma_s^2 \right\} \\ & - \frac{1}{2} \sum_{i=1}^N \sum_{t=1}^{T_i} \left\{ (y_{it} - x_{it}\beta + z_{it}\delta)^2 / \sigma_s^2 \right\} \\ & - \sum_{i=1}^N \sum_{t=1}^{T_i} \left\{ \ln \Phi(d_{it}) - \ln \Phi(d_{it}^*) \right\} \end{aligned} \quad (A2.13)$$

$$\text{where } d_{it} = z_{it} \delta / (\gamma \sigma_S^2)^{1/2} \quad (\text{A2.14})$$

$$d_{it}^* = \mu_{it}^* / [\gamma(1-\gamma)\sigma_S^2]^{1/2} \quad (\text{A2.15})$$

$$\mu_{it}^* = (1-\gamma)z_{it}\delta - \gamma(y_{it} - x_{it}\beta) \quad (\text{A2.16})$$

$$\sigma_* = [\gamma(1-\gamma)\sigma_S^2]^{1/2} \quad (\text{A2.17})$$

$$\text{and } \theta = (\beta', \delta', \sigma_S^2, \gamma)'$$

The partial derivatives of the logarithm of the likelihood function (A2.13) with respect to the parameters, β , δ , σ_S^2 , and γ , are given by

$$\begin{aligned} \frac{\partial L^*}{\partial \beta} &= \sum_{i=1}^N \sum_{t=1}^{T_i} \left\{ \frac{(y_{it} - x_{it}\beta + z_{it}\delta)}{\sigma_S^2} + \frac{\phi(d_{it}^*)}{\Phi(d_{it}^*)} \frac{\gamma}{\sigma_*} \right\} x'_{it} \\ \frac{\partial L^*}{\partial \delta} &= - \sum_{i=1}^N \sum_{t=1}^{T_i} \left\{ \frac{(y_{it} - x_{it}\beta + z_{it}\delta)}{\sigma_S^2} + \left[\frac{\phi(d_{it})}{\Phi(d_{it})} \cdot \frac{1}{(\gamma \sigma_S^2)^{1/2}} - \frac{\phi(d_{it}^*)}{\Phi(d_{it}^*)} \cdot \frac{(1-\gamma)}{\sigma_*} \right] \right\} z'_{it} \\ \frac{\partial L^*}{\partial \sigma_S^2} &= -\frac{1}{2} \left(\frac{1}{\sigma_S^2} \right) \left\{ \left(\sum_{i=1}^N T_i \right) - \sum_{i=1}^N \sum_{t=1}^{T_i} \left[\frac{\phi(d_{it})}{\Phi(d_{it})} d_{it} - \frac{\phi(d_{it}^*)}{\Phi(d_{it}^*)} d_{it}^* \right] \right. \\ &\quad \left. - \sum_{i=1}^N \sum_{t=1}^{T_i} \frac{(y_{it} - x_{it}\beta + z_{it}\delta)}{\sigma_S^2} \right\} \\ \frac{\partial L^*}{\partial \gamma} &= \sum_{i=1}^N \sum_{t=1}^{T_i} \left\{ \frac{\phi(d_{it})}{\Phi(d_{it})} \frac{d_{it}}{2\gamma} + \frac{\phi(d_{it}^*)}{\Phi(d_{it}^*)} \left[\frac{y_{it} - x_{it}\beta + z_{it}\delta}{\sigma_*} + \frac{d_{it}^*(1-2\gamma)}{2\gamma(1-\gamma)\sigma_*^2} \right] \right\}. \end{aligned}$$

Appendix 3: Electricity Data

Important details on the sources and definitions of data used in the study in Chapter 6 are discussed below.

Missing data:

Two plants began operation during the sample period. These were Loy Yang A and Bayswater, which were commissioned in 1983/84 and 1985/86, respectively.

Data on labour was not available for 1988/89 in the four Victorian plants because of a change over from local maintenance staff to a maintenance group which moved from plant to plant. This change made labour figures difficult to derive.

The NSW data were collected by Michael Plumb from annual reports for use in an undergraduate project at the University of Sydney, and was kindly supplied for this analysis. The data that were supplied did not include observations for 1981/82. Hence this year are not considered for the six NSW plants in this study.

Production data were not available for Bunbury in 1981/82. These were the only observations missing from the three WA plants.

Derived data:

Data on number of employees were not available from Victoria, but data on wages and salaries of employees were available. Hence the number of employees in Victorian plants was derived using an implicit wage series obtained from the WA data. The WA expenditure on employees was divided by the number of WA employees to obtain an implicit wage series. It was then assumed that similar wages would be paid in Victoria so that the expenditure on employees in Victorian plants could be divided by this wage series to obtain a measure of the number of employees in Victorian plants.

The average age of generating units is a weighted average of the ages of the units at a plant, where the weights are the name-plate capacities of the units.

The output measure is electricity sent out. Capacity factor is calculated using this figure, not the electricity generated.

The measure of total fuel usage in terrajoules was derived by converting each of the fuels into their energy equivalents. The energy contents assumed for the various fuels are listed in Table A3.1.

Table A3.1
Energy Contents of Fuels (in megajoules/kilogram) in 13 Power Plants in Australia

Plant	Coal	Oil	Briquette s
Loy Yang A	8.6		22.1
Hazelwood	8.8		22.1
Yallourn W	7.4	43.7	
Morwell	8.8	43.7	
Bayswater	23.4	45.9	
Eraring	23.9	45.9	
Liddell	22.2	46.4	
Munmorah	26.4	45.9	
Vales Point B	23.8	43.7	
Wallerang C	25.0	45.9	
Bunbury	20.0	43.0	
Muja	20.0	43.0	
Kwinana	19.9	43.0	

Table A3.2
Data From 13 Australian Coal-Fired Power Plants

Year	Plant	cap (mw)	output (kwh)	coal (000t)	gas (tj)	oil (000t)	briqts (000t)	lab (per)	tot-fuel (tj)	a/age (yrs)	a/size (mw)	c/fac (%)
83/84	loy	500	376	404.8	0	0.0	98.4	41	5656	0	500	8.58
84/85	loy	1000	4057	5007.4	0	0.0	87.8	458	45004	1	500	46.31
85/86	loy	1500	6160	7547.0	0	0.0	87.0	569	66827	1	500	46.88
86/87	loy	2000	9092	10857.9	0	0.0	146.1	616	96607	2	500	51.89
87/88	loy	2000	12307	13725.3	0	0.0	123.3	802	120763	2	500	70.25
88/89	loy	2000	13616	16859.0	0	0.0	87.0	-9	146910	3	500	77.72
89/90	loy	2000	14920	18028.0	0	0.0	60.0	695	156367	4	500	85.16

90/91	loy	2000	14873	18030.0	109	0.0	47.5	534	156217	5	500	84.89
81/82	haz	1600	8174	11646.8	0	0.0	216.3	1566	107272	14	200	58.32
82/83	haz	1600	6697	9566.3	0	0.0	146.4	1585	87419	15	200	47.78
83/84	haz	1600	6829	9368.4	0	0.0	136.8	1787	85465	16	200	48.72
84/85	haz	1600	7565	11031.3	0	0.0	80.5	1634	98854	17	200	53.97
85/86	haz	1600	6984	10034.6	0	0.0	74.8	1528	89958	18	200	49.83
86/87	haz	1600	7272	10354.3	0	0.0	81.1	1266	92910	19	200	51.88
87/88	haz	1600	7285	10306.0	0	0.0	42.7	1191	91636	20	200	51.98
88/89	haz	1600	8397	12361.0	0	0.0	63.4	-9	110178	21	200	59.91
89/90	haz	1600	6995	10468.0	0	0.0	76.0	1675	93798	22	200	49.91
90/91	haz	1600	8445	12435.6	9.45	0.0	38.6	1127	110296	23	200	60.25
81/82	yal	1450	7608	11816.0	0	90.3	0.0	988	91385	4	362.5	59.90
82/83	yal	1450	8882	14440.8	0	45.0	0.0	1059	108828	5	362.5	69.93
83/84	yal	1450	8248	13487.1	0	28.4	0.0	1103	101046	6	362.5	64.93
84/85	yal	1450	8375	13897.3	0	17.7	0.0	1096	103614	7	362.5	65.93
85/86	yal	1450	7551	11548.6	0	20.0	0.0	1084	86334	8	362.5	59.45
86/87	yal	1450	8460	13891.2	0	13.6	0.0	986	103389	9	362.5	66.60
87/88	yal	1450	8006	13140.6	0	13.7	0.0	901	97839	10	362.5	63.03
88/89	yal	1450	8570	14016.6	0	24.7	0.0	-9	104802	11	362.5	67.47
89/90	yal	1450	8463	13200.0	0	29.6	0.0	923	98974	12	362.5	66.63
90/91	yal	1450	9479	14690.1	0	17.0	0.0	648	109450	13	362.5	74.63
81/82	mor	170	1010	2092.9	0	1.2	0.0	476	18470	22	34	67.82
82/83	mor	170	1092	2115.0	0	1.3	0.0	467	18669	23	34	73.33
83/84	mor	170	1059	2073.8	0	1.2	0.0	452	18302	24	34	71.11
84/85	mor	170	868	1821.0	0	1.4	0.0	466	16086	25	34	58.29
85/86	mor	170	805	1810.2	0	1.8	0.0	435	16008	26	34	54.06
86/87	mor	170	762	1721.5	0	1.1	0.0	399	15197	27	34	51.17
87/88	mor	170	742	1757.6	0	1.3	0.0	396	15524	28	34	49.83
88/89	mor	170	835	1816.1	0	1.0	0.0	-9	16025	29	34	56.07
89/90	mor	170	531	1371.6	0	1.3	0.0	387	12127	30	34	35.66
90/91	mor	170	462	1203.6	0	1.2	0.0	344	10644	31	34	31.02
85/86	bay	1320	5562	2792.0	0	6.8	0.0	400	65645	0	660	48.10
86/87	bay	2640	11451	5175.0	0	8.5	0.0	519	121485	1	660	49.51
87/88	bay	2640	10288	4540.0	0	8.5	0.0	561	106626	1	660	44.49
88/89	bay	2640	12593	5640.0	0	6.7	0.0	599	132284	2	660	54.45
89/90	bay	2640	14182	6116.0	0	3.3	0.0	603	143266	3	660	61.32
90/91	bay	2640	15406	6603.0	0	3.5	0.0	556	154671	4	660	66.62
82/83	era	1320	5817	2526.0	0	6.5	0.0	390	60670	0	660	50.31

83/84	era	1980	8337	3586.0	0	9.7	0.0	518	86151	1	660	48.07
84/85	era	2640	10089	4580.0	0	10.6	0.0	560	109949	1	660	43.63
85/86	era	2640	8931	4147.0	0	10.1	0.0	641	99577	2	660	38.62
86/87	era	2640	9680	4542.0	0	8.4	0.0	672	108939	3	660	41.86
87/88	era	2640	10284	4568.0	0	8.4	0.0	669	109561	4	660	44.47
88/89	era	2640	11429	4980.0	0	8.0	0.0	694	119389	5	660	49.42
89/90	era	2640	11230	4908.0	0	6.3	0.0	644	117590	6	660	48.56
90/91	era	2640	12539	5279.0	0	4.0	0.0	583	126352	7	660	54.22
82/83	lid	2000	8551	4448.0	0	6.5	0.0	662	99047	10	500	48.81
83/84	lid	2000	9966	5096.0	0	13.4	0.0	683	113753	11	500	56.88
84/85	lid	2000	9980	5133.0	0	14.2	0.0	645	114611	12	500	56.96
85/86	lid	2000	6948	3601.0	0	18.0	0.0	653	80777	13	500	39.66
86/87	lid	2000	6214	3146.0	0	22.5	0.0	652	70885	14	500	35.47
87/88	lid	2000	6660	3433.0	0	13.9	0.0	645	76858	15	500	38.01
88/89	lid	2000	5673	2883.0	0	13.2	0.0	590	64615	16	500	32.38
89/90	lid	2000	6177	3150.0	0	10.9	0.0	571	70436	17	500	35.26
90/91	lid	2000	5190	2569.0	0	7.8	0.0	554	57394	18	500	29.62
82/83	mun	1200	5482	2538.0	0	5.8	0.0	562	67269	13	300	52.15
83/84	mun	1200	4927	2268.0	0	4.3	0.0	565	60073	14	300	46.87
84/85	mun	1200	3451	1622.0	0	4.5	0.0	570	43027	15	300	32.83
85/86	mun	1200	3937	1773.0	0	2.9	0.0	585	46940	16	300	37.45
86/87	mun	1200	3509	1582.0	0	4.3	0.0	614	41962	17	300	33.38
87/88	mun	1200	3094	1385.0	0	4.9	0.0	609	36789	18	300	29.43
88/89	mun	1200	3216	1502.0	0	13.1	0.0	576	40254	19	300	30.59
89/90	mun	1200	3125	1406.0	0	3.8	0.0	589	37293	20	300	29.73
90/91	mun	1200	2652	1507.0	0	4.6	0.0	452	39996	21	300	25.23
82/83	val	1320	7481	3173.0	0	9.3	0.0	849	75924	3	660	64.70
83/84	val	1320	8176	3254.0	0	6.1	0.0	873	77712	4	660	70.71
84/85	val	1320	7572	3176.0	0	11.5	0.0	854	76091	5	660	65.48
85/86	val	1320	6814	2907.0	0	11.3	0.0	884	69680	6	660	58.93
86/87	val	1320	5945	2618.0	0	8.3	0.0	852	62671	7	660	51.41
87/88	val	1320	6118	2745.0	0	13.6	0.0	800	65925	8	660	52.91
88/89	val	1320	5672	2828.0	0	13.7	0.0	696	67905	9	660	49.05
89/90	val	1320	6589	2777.0	0	4.6	0.0	449	66294	10	660	56.98
90/91	val	1320	6309	2662.0	0	4.6	0.0	417	63557	11	660	54.56
82/83	wal	1000	3271	1714.0	0	8.4	0.0	591	43236	4	500	37.34
83/84	wal	1000	2703	1297.0	0	4.3	0.0	596	32622	5	500	30.86
84/85	wal	1000	3545	1595.0	0	5.5	0.0	573	40127	6	500	40.47

85/86	wal	1000	5238	2311.0	0	3.9	0.0	586	57954	7	500	59.79
86/87	wal	1000	2458	1149.0	0	6.5	0.0	585	29023	8	500	28.06
87/88	wal	1000	4479	1888.0	0	4.7	0.0	593	47416	9	500	51.13
88/89	wal	1000	4171	1883.0	0	4.3	0.0	520	47272	10	500	47.61
89/90	wal	1000	4259	1901.0	0	3.8	0.0	496	47699	11	500	48.62
90/91	wal	1000	2750	1254.0	0	3.4	0.0	451	31506	12	500	31.39
81/82	bun	120	-9	323.8	0	1.9	0.0	182	6557	23	30	na
82/83	bun	120	401	282.5	0	2.4	0.0	179	5755	24	30	38.19
83/84	bun	120	374	269.9	0	2.6	0.0	171	5508	25	30	35.64
84/85	bun	120	155	115.3	0	2.4	0.0	195	2409	26	30	14.78
85/86	bun	120	99	80.3	0	1.5	0.0	144	1671	27	30	9.47
86/87	bun	120	94	74.2	0	1.5	0.0	138	1548	28	30	9.01
87/88	bun	120	65	60.6	0	1.6	0.0	132	1282	29	30	6.19
88/89	bun	120	177	136.9	0	1.8	0.0	134	2816	30	30	16.86
89/90	bun	120	165	127.4	0	1.5	0.0	136	2612	31	30	15.76
90/91	bun	120	129	103.3	0	1.4	0.0	136	2123	32	30	12.30
81/82	muja	640	2844	1764.4	0	6.5	0.0	464	35569	6	107	50.74
82/83	muja	640	3233	2000.1	0	5.7	0.0	525	40248	7	107	57.67
83/84	muja	640	3113	1919.8	0	6.5	0.0	562	38677	8	107	55.54
84/85	muja	840	3973	2353.1	0	6.2	0.0	579	47327	9	107	53.99
85/86	muja	1040	3723	2242.2	0	8.8	0.0	732	45221	7	120	40.87
86/87	muja	1040	3563	1981.7	0	11.5	0.0	700	40127	6	130	39.12
87/88	muja	1040	3698	2251.9	0	6.6	0.0	668	45323	7	130	40.60
88/89	muja	1040	4828	2865.5	0	10.0	0.0	704	57741	8	130	53.00
89/90	muja	1040	5549	3210.0	0	10.9	0.0	741	64667	9	130	60.92
90/91	muja	1040	6111	3471.9	0	9.2	0.0	736	69831	10	130	67.08
81/82	kwin	880	1502	684.7	0	73.6	0.0	377	16793	9	147	19.49
82/83	kwin	880	1336	704.1	0	61.5	0.0	382	16656	10	147	17.34
83/84	kwin	880	1698	940.2	537	27.3	0.0	424	20423	11	147	22.03
84/85	kwin	880	1817	852.7	3220	18.9	0.0	432	20999	12	147	23.58
85/86	kwin	880	2443	63.8	27060	8.5	0.0	456	28694	13	147	31.69
86/87	kwin	880	3025	112.9	33137	3.7	0.0	439	35544	14	147	39.25
87/88	kwin	880	3512	84.4	40610	1.1	0.0	458	42336	15	147	45.56
88/89	kwin	880	3115	230.5	32288	4.7	0.0	481	37075	16	147	40.41
89/90	kwin	880	3076	325.5	29457	6.3	0.0	473	36205	17	147	39.91
90/91	kwin	880	2913	723.2	20140	1.1	0.0	503	34581	18	147	37.80

Appendix 4: FRONTIER Programmer's Guide

The FRONT41.000 File

The start-up file FRONT41.000 is listed in Table A4.1. Ten values may be altered in FRONT41.000. A brief description of each value is provided below.

Table A4.1

The Start-up File for the FRONTIER Program: FRONT41.000

KEY VALUES USED IN FRONTIER PROGRAM (VERSION 4.1)

NUMBER: DESCRIPTION:

5	IPRINT - PRINT INFO EVERY "N" ITERATIONS, 0=DO NOT PRINT
1	INDIC - USED IN UNIDIMENSIONAL SEARCH PROCEDURE - SEE BELOW
0.00001	TOL - CONVERGENCE TOLERANCE (PROPORTIONAL)
0.001	TOL2 - TOLERANCE USED IN UNI-DIMENSIONAL SEARCH PROCEDURE
1.0D+16	BIGNUM - USED TO SET BOUNDS ON DEN & DIST
0.00001	STEP1 - SIZE OF 1ST STEP IN SEARCH PROCEDURE
1	IGRID2 - 1=DOUBLE ACCURACY GRID SEARCH, 0=SINGLE
0.1	GRIDNO - STEPS TAKEN IN SINGLE ACCURACY GRID SEARCH ON GAMMA
100	MAXIT - MAXIMUM NUMBER OF ITERATIONS PERMITTED
1	ITE - 1=PRINT ALL TE ESTIMATES, 0=PRINT ONLY MEAN TE

THE NUMBERS IN THIS FILE ARE READ BY THE FRONTIER PROGRAM WHEN IT BEGINS EXECUTION. YOU MAY CHANGE THE NUMBERS IN THIS FILE IF YOU WISH. IT IS ADVISED THAT A BACKUP OF THIS FILE IS MADE PRIOR TO ALTERATION.

FOR MORE INFORMATION ON THESE VARIABLES SEE WORKING PAPER: COELLI (1994) FROM THE UNIVERSITY OF NEW ENGLAND, ARMIDALE, NSW, 2351, AUSTRALIA.

INDIC VALUES:

indic=2 says do not scale step length in unidimensional search

indic=1 says scale (to length of last step) only if last step was smaller

indic= any other number says scale (to length of last step)

- 1) IPRINT - specifies how often information on the likelihood function value and the vector of parameter estimates should be recorded during the iterative process. It is initially set to 5, hence information is printed every 5 iterations. It can be set to

any non-negative integer value. A 0 will result in no reporting of intermediate information.

- 2) INDIC - relates to the Coggin uni-dimensional search which is conducted before each iteration to determine the optimal step length. It may be used as follows: indic=2 says do not scale step length in uni-dimensional search; indic=1 says scale (to length of last step) only if last step was smaller; and indic=any other number says scale (to length of last step) For more information see Himmelblau (1972).
- 3) TOL - sets the convergence tolerance on the iterative process. If this value is say set to 0.00001 then the iterative procedure would terminate when the proportional change in the log-likelihood function and in each of the estimated parameters are all less than 0.00001.
- 4) TOL2 - sets the required tolerance on the Coggin uni-dimensional search done each iteration to determine the step length. For more information see Himmelblau (1972).
- 5) BIGNUM - bounds the size of the largest number that the program should deal with. Its primary use is to place bounds upon what the smallest number can be in the subroutines DEN (which evaluates the standard normal density function) and DIS (which evaluates the standard normal distribution function). Errors with numerical underflows and overflows were the problems most frequently encountered by people attempting to install earlier versions of this program on various mainframe computers. This number has been set to 1.0e+16 for the IBM PC. If you plan to mount this program on a mainframe computer it is advised that you consult computer support staff on the correct setting of this number. It generally would be safe to leave it as it is, however, greater precision may be gained if larger numbers are permitted.
- 6) STEP1 - sets the size of the first step in the iterative process. This should be set carefully as too large a value may result in the program 'stepping' right out of the sensible parameter space.
- 7) IGRID2 - a flag which if set to 1 will cause the grid search to complete a second phase grid search around the estimate obtained in the first phase of the grid search.

If set to zero only the first phase of the grid search will be conducted. For more information refer to the description of the grid search in Section 3.

- 8) GRIDNO - sets the width of the steps taken in the grid search between zero and one on the γ parameter. For more information refer to the description of the grid search in Section 3.
- 9) MAXIT - sets the maximum number of iterations that will be conducted. This is especially a useful option when batch files are written for monte carlo simulation.
- 10) ITE - specifies whether individual efficiency estimates should be listed in the output file. A value of 1 will cause them to be listed, and a 0 will suppress them.

Subroutine and Function Descriptions

SUBROUTINES:

- 1) EXEC: This is the main calling program. It firstly reads the start-up file FRONT2.000 before calling INFO.
- 2) INFO: This subroutine reads instructions either from a file or from the terminal then reads the data file. It then calls MINI.
- 3) MINI: This is the main subroutine of the program. It firstly calls GRID to do the grid search (assuming starting values are not specified by the user). MINI then conducts the main iterative loop of FRONTIER, calling SEARCH, ETA and CONVRG repeatedly until the convergence criteria are satisfied (or the maximum number of iterations is achieved). The Davidon-Fletcher-Powell method is used.
- 4) RESULT: Sends all final results to the output file. These include parameter estimates, approximate standard errors, t-ratios, and the individual and mean technical efficiency estimates.
- 5) GRID: Does a grid search over γ .
- 6) SEARCH: Performs a uni-dimensional search to determine the optimal step length for the next iteration. The Coggin method is used (see Himmelblau, 1972).

- 7) ETA: This subroutine updates the direction matrix according to the Davidon-Fletcher-Powell method at each iteration. For more information refer to Himmelblau (1972).
- 8) CONVRG: Tests the convergence critereon at the end of each iteration. If the proportion change in the log-likelihood function and each of the parameters is no greater than the value of TOL (initially set to 0.00001) the iterative process will terminate.
- 9) FUN1: Calculates the negative of the log-likelihood function (LLF) of Model 1. Note that FRONTIER minimizes the negative of the LLF which is equivalent to maximizing the LLF.
- 10) DER1: Calculates the first partial derivatives of the negative of the LLF of Model 1.
- 11) FUN2: Calculates the negative of the log-likelihood function (LLF) of Model 2.
- 12) DER2: Calculates the first partial derivatives of the negative of the LLF of Model 2.
- 13) CHECK: Ensures that the estimated parameters do not venture outside the theoretical bounds (i.e. $0 < \gamma < 1$, $\sigma^2 > 0$ and $-2\sigma_U < \mu < 2\sigma_U$).
- 14) OLS: Calculates the Ordinary Least Squares estimates of the model to be used as starting values. It also calculates OLS standard errors which are presented in the final output.
- 15) INVERT: Inverts a given matrix.

FUNCTIONS:

- 1) DEN: Evaluates the density function of a standard normal random variable.
- 2) DIS: Evaluates the distribution function of a standard normal random variable.

Appendix 5: FRONTIER Program Listing

```

program exec
c   FRONTIER version 4.1 by Tim Coelli.
c   This program uses the Davidon-Fletcher-Powell algorithm to
c   estimate two forms of the stochastic frontier production function.
c   The first is the error components model described in Battese and
c   Coelli (1992) in the JPA, and the second is the TE effects model
c   described in Battese and Coelli (1993), UNE Working Paper #69.
c   A large proportion of the search, convrg, mini and eta
c   subroutines are taken from the book: Himmelblau (1972, appendix b).
c   The remainder of the program is the work of Tim Coelli.
c   Any person is welcome to copy and use this program free of charge.
c   If you find the program useful, a contribution of $150 US
c   to help defray some of the author's costs would be appreciated, but
c   is in no way obligatory.
c   Please note that the author takes no responsibility for any
c   inconvenience caused by undetected errors. If an error is
c   detected the author would appreciate being informed. He may be
c   contacted via email at tcoelli@metz.une.edu.au .
c   Postal address: Tim Coelli, Department of Econometrics,
c   University of New England, Armidale, NSW 2351, Australia,
c   last update = 17/October/1994
implicit double precision (a-h,o-z)
character*12 koutf,kdatf,kinf
common/eight/koutf,kdatf,kinf
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
common/one/fx,fy,nm,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
open(unit=70,file='front41.000',status='old')
read(70,*)
read(70,*)
read(70,*) iprint
read(70,*) indic
read(70,*) tol
read(70,*) tol2
read(70,*) bignum
read(70,*) step1
read(70,*) igrid2
read(70,*) gridno
read(70,*) maxit
read(70,*) ite
close(70)
nfunct=0
ndrv=0
call info
stop
end

subroutine mini(yy,xx,mm,sv)
c   contains the main loop of this iterative program.
implicit double precision (a-h,o-z)
character*12 koutf,kdatf,kinf
common/eight/koutf,kdatf,kinf
common/one/fx,fy,nm,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit

```

```

dimension yy(nn,nt),xx(nn,nt,nr),mm(nn),sv(n)
dimension ob(:),gb(:),obse(:),x(:),y(:),s(:)
dimension h(:,:),delx(:),delg(:),gx(:),gy(:)
allocatable :: ob,gb,obse,x,y,s,h,delx,delg,gx,gy
allocate(ob(n),gb(n),obse(n),x(n),y(n),s(n))
allocate(h(n,n),delx(n),delg(n),gx(n),gy(n))
open(unit=70,file=koutf,status='unknown')
do 98 i=1,n
  gx(i)=0.0
  gy(i)=0.0
98 continue
  call ols(ob,obse,yy,xx)
  if (igrd.eq.1) then
    write(6,*) 'doing grid search...'
    call grid(x,y,yy,xx,ob,gb)
  else
    do 131 i=1,n
      y(i)=sv(i)
      x(i)=sv(i)
131 continue
    if (im.eq.1) call fun1(x,fx,yy,xx)
    if (im.eq.2) call fun2(x,fx,yy,xx)
    fy=fx
    end if
    call result(yy,xx,mm,h,y,sv,ob,obse,gb,1)
    write(6,*) 'DFP iterative process underway - please wait...'
    iter=0
    if (im.eq.1) call der1(x,gx,yy,xx)
    if (im.eq.2) call der2(x,gx,yy,xx)
    write(70,*)
    write(70,*)
    write(70,301) iter,nfunct,-fy
    nc=1
305 write(70,302) (y(i),i=nc,min(n,nc+4))
    nc=nc+5
    if(nc.le.n) goto 305
    if (maxit.eq.0) goto 70
5   do 20 i=1,n
      do 10 j=1,n
10    h(i,j)=0.0
20    h(i,i)=1.0
      if(iprint.ne.0) write(70,2100)
2100 format(' gradient step')
      do 30 i=1,n
30    s(i)=-gx(i)
40    call search(x,y,s,gx,delx,yy,xx)
      iter=iter+1
      if (iter.ge.maxit) then
        write(70,*) 'maximum number of iterations reached'
        goto 70
      endif
7   if(fy.gt.fx) goto 5
      if (im.eq.1) call der1(y,gy,yy,xx)
      if (im.eq.2) call der2(y,gy,yy,xx)
      call convrg(ipass,x,y)
      if (ipass.eq.1.) goto 70
      if (iprint.ne.0) then
        printcon=float(iter)/float(iprint)-float(iter/iprint)

```

```

        if (printcon.eq.0.0) then
        write(70,301) iter,nfunct,-fy
        nc=1
304  write(70,302) (y(i),i=nc,min(n,nc+4))
        nc=nc+5
        if(nc.le.n) goto 304
        endif
        endif
        do 50 i=1,n
        delg(i)=gy(i)-gx(i)
        delx(i)=y(i)-x(i)
        gx(i)=gy(i)
50  x(i)=y(i)
        fx=fy
        call eta(h,delx,delg,gx)
        do 60 i=1,n
        s(i)=0.0
        do 60 j=1,n
60  s(i)=s(i)-h(i,j)*gy(j)
        goto 40
70  continue
        write(70,301) iter,nfunct,-fy
        nc=1
303  write(70,302) (y(i),i=nc,min(n,nc+4))
        nc=nc+5
        if(nc.le.n) goto 303
301  format(' iteration = ',i5,' func evals = ',i7,' llf = ',e16.8)
302  format(4x,5e15.8)
        call result(yy,xx,mm,h,y,sv,ob,obse,gb,2)
        deallocate(ob,gb,obse,x,y,s,h,delx,delg,gx,gy)
        close(70)
        return
        end

```

```

subroutine convrg(ipass,x,y)
c  tests the convergence criterion.
c  the program is halted when the proportional change in the log-
c  likelihood and in each of the parameters is no greater than
c  a specified tolerance.
implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
dimension x(n),y(n)
xtol=tol
ftol=tol
if(dabs(fx).le.ftol) goto 10
if(dabs((fx-fy)/fx).gt.ftol) goto 60
goto 20
10 if(dabs(fx-fy).gt.ftol) goto 60
20 do 40 i=1,n
    if(dabs(x(i)).le.xtol) goto 30
    if(dabs((x(i)-y(i))/x(i)).gt.xtol) goto 60
    goto 40
30 if(dabs(x(i)-y(i)).gt.xtol) goto 60
40 continue
    ipass=1
    return

```

```

60  ipass=2
    return
    end

    subroutine eta(h,delx,delg,gx)
c    calculates the direction matrix (p).
    implicit double precision (a-h,o-z)
    common/three/n,nfunct,ndrv,iter,indic,iprint,igrd,maxit
    common/five/tol,tol2,bignum,step1,gridno,igrd2,ite
    dimension h(n,n),delx(n),delg(n),gx(n)
    dimension hdg(:),dgh(:),hgx(:)
    allocatable :: hdg,dgh,hgx
    allocate(hdg(n),dgh(n),hgx(n))
    dxdg=0.0
    dghdg=0.0
    do 20 i=1,n
        hdg(i)=0.0
        dgh(i)=0.0
        do 10 j=1,n
            hdg(i)=hdg(i)-h(i,j)*delg(j)
10        dgh(i)=dgh(i)+delg(j)*h(j,i)
            dxdg=dxdg+delx(i)*delg(i)
20        dghdg=dghdg+dgh(i)*delg(i)
            do 30 i=1,n
                do 30 j=1,n
30            h(i,j)=h(i,j)+delx(i)*delx(j)/dxdg+hdg(i)*dgh(j)/dghdg
                do 117 i=1,n
117            h(i,i)=dabs(h(i,i))
                do 132 i=1,n
                    hgx(i)=0.0
                    do 132 j=1,n
                        hgx(i)=hgx(i)+h(i,j)*gx(j)
132            continue
                    hgxx=0.
                    gxx=0.
                    do 133 i=1,n
                        hgxx=hgxx+hgx(i)**2
                        gxx=gxx+gx(i)**2
133            continue
                    c=0.
                    do 134 i=1,n
                        c=c+hgx(i)*gx(i)
134            continue
                    c=c/(hgxx*gxx)**0.5
                    if(dabs(c).lt.1.0/bignum) then
                        write(6,*) 'ill-conditioned eta'
                        do 136 i=1,n
                            do 137 j=1,n
137            h(i,j)=0.0
136            h(i,i)=delx(i)/gx(i)
                        endif
                    deallocate(hdg,dgh,hgx)
                    return
                    end

    subroutine search(x,y,s,gx,delx,yy,xx)
c    unidimensional search (coggin) to determine optimal step length
c    determines the step length (t) using a unidimensional search.

```

```

implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
dimension x(n),y(n),s(n),gx(n),delx(n)
dimension yy(nn,nt),xx(nn,nt,nr)
iexit=0
ntol=0
ftol=tol2
ftol2=ftol/100.0
fa=fx
fb=fx
fc=fx
da=0.0
db=0.0
dc=0.0
k=-2
m=0
step=step1
d=step
if(indic.eq.2.or.iter.eq.0) goto 1
dxnorm=0.0
snorm=0.0
do 102 i=1,n
dxnorm=dxnorm+delx(i)*delx(i)
102 snorm=snorm+s(i)*s(i)
if(indic.eq.1.and.dxnorm.ge.snorm) goto 1
ratio=dxnorm/snorm
step=dsqrt(ratio)
d=step
1 do 2 i=1,n
2 y(i)=x(i)+d*s(i)
if (im.eq.1) call fun1(y,f,yy,xx)
if (im.eq.2) call fun2(y,f,yy,xx)
k=k+1
if(f-fa) 5,3,6
3 do 4 i=1,n
4 y(i)=x(i)+da*s(i)
fy=fy
if(iprint.ne.0) write(70,2100)
2100 format(' search failed. fn val indep of search direction')
goto 326
5 fc=fb
fb=fa
fa=f
dc=db
db=da
da=d
d=2.0*d+step
goto 1
6 if(k) 7,8,9
7 fb=f
db=d
d=-d
step=-step
goto 1
8 fc=fb
fb=fa

```

```

fa=f
dc=db
db=da
da=d
goto 21
9 dc=db
db=da
da=d
fc=fb
fb=fa
fa=f
10 d=0.5*(da+db)
do 11 i=1,n
11 y(i)=x(i)+d*s(i)
if (im.eq.1) call fun1(y,f,yy,xx)
if (im.eq.2) call fun2(y,f,yy,xx)
12 if((dc-d)*(d-db)) 15,13,18
13 do 14 i=1,n
14 y(i)=x(i)+db*s(i)
fy=fb
if(iexit.eq.1) goto 32
if(iprint.ne.0) write(70,2500)
2500 format(' search failed. loc of min limited by rounding')
goto 325
15 if(f-fb) 16,13,17
16 fc=fb
fb=f
dc=db
db=d
goto 21
17 fa=f
da=d
goto 21
18 if(f-fb) 19,13,20
19 fa=fb
fb=f
da=db
db=d
goto 21
20 fc=f
dc=d
21 a=fa*(db-dc)+fb*(dc-da)+fc*(da-db)
if(a) 22,30,22
22 d=0.5*((db*db-dc*dc)*fa+(dc*dc-da*da)*fb+(da*da-db*db)*fc)/a
if((da-d)*(d-dc)) 13,13,23
23 do 24 i=1,n
24 y(i)=x(i)+d*s(i)
if (im.eq.1) call fun1(y,f,yy,xx)
if (im.eq.2) call fun2(y,f,yy,xx)
if(dabs(fb)-ftol2) 25,25,26
25 a=1.0
goto 27
26 a=1.0/fb
27 if((dabs(fb-f)*a)-ftol) 28,28,12
28 iexit=1
if(f-fb) 29,13,13
29 fy=f
goto 32

```



```

30 if(m) 31,31,13
31 m=m+1
   goto 10
32 do 99 i=1,n
   if(y(i).ne.x(i)) goto 325
99 continue
   goto 33
325 if(ntol.ne.0.and.iprint.eq.1) write(70,3000) ntol
3000 format(1x,'tolerance reduced',i1,'time(s)')
326 if(fy.lt.fx) return
   do 101 i=1,n
   if(s(i).ne.-gx(i)) return
101 continue
   write(70,5000)
5000 format(' search failed on gradient step, termination')
   return
33 if(ntol.eq.5) goto 34
   iexit=0
   ntol=ntol+1
   ftol=ftol/10.
   goto 12
34 if(iprint.ne.0) write(70,2000)
2000 format(' pt better than entering pt cannot be found')
   return
   end

subroutine check(b,xx)
c   checks if params are out of bounds & adjusts if required.
implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
dimension b(n),xx(nn,nt,nr)
n1=nr+1
n2=nr+2
bi=dsqrt(bignum)
if(b(n1).le.0.0) b(n1)=1.0/bi
if(b(n2).le.1.0/bi) b(n2)=1.0/bi
if(b(n2).ge.1.0-1.0/bi) b(n2)=1.0-1.0/bi
bound=2.*dsqrt(b(n1)*b(n2))
if((im.eq.1).and.(nmu.eq.1)) then
n3=nr+3
if(b(n3).gt.bound) b(n3)=bound
if(b(n3).lt.-bound) b(n3)=-bound
endif
return
end

subroutine fun1(b,a,yy,xx)
c   calculates the negative of the log-likelihood function of the
c   error components model.
implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
data pi/3.1415926/
dimension b(n),yy(nn,nt),xx(nn,nt,nr)
call check(b,xx)
a=0.0

```

```

f=dfloat(nn)
fnt=dfloat(nt)
ftot=dfloat(nob)
s2=b(nb+1)
g=b(nb+2)
u=0.0
e=0.0
if (nmueq.1) then
  u=b(nb+3)
  if (netaeq.1) e=b(nb+4)
  else
    if (netaeq.1) e=b(nb+3)
  endif
  sc=1.
  if (ipceq.2) sc=-1.
  a=0.5*ftot*(dlog(2.0*pi)+dlog(s2))
  a=a+0.5*(ftot-f)*dlog(1.0-g)
  z=u/(s2*g)**0.5
  a=a+f*dlog(dis(z))
  a=a+0.5*f*z**2
  a2=0.0
  do 132 i=1,nn
    epr=0.0
    do 103 l=1,nt
      if (xx(i,l,1).ne.0.0) then
        ee=yy(i,l)
        do 102 j=1,nb
          ee=ee-b(j)*xx(i,l,j)
102      continue
          epr=epr+ee*dexp(-e*(dfloat(l)-fnt))
        end if
103      continue
          epe=0.0
          do 101 l=1,nt
            if (xx(i,l,1).ne.0.0) epe=epe+dexp(-2.0*e*(dfloat(l)-fnt))
101          continue
            zi=(u*(1.0-g)-sc*g*epr)/(g*(1.0-g)*s2*(1.0+(epe-1.0)*g))**0.5
            a=a+0.5*dlog(1.0+(epe-1.0)*g)
            a=a-dlog(dis(zi))
            do 133 l=1,nt
              if (xx(i,l,1).ne.0.0) then
                ee=yy(i,l)
                do 134 j=1,nb
                  ee=ee-b(j)*xx(i,l,j)
134              continue
                  a2=a2+ee**2
                end if
133              continue
                  a=a-0.5*zi**2
132              continue
                  a=a+0.5*a2/((1.0-g)*s2)
            nfunct=nfunct+1
            return
          end
    end

subroutine der1(b,gx,yy,xx)
c    calculates the first-order partial derivatives of the negative
c    of the log-likelihood function of the error components model.

```

```

implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
dimension b(n),gx(n),yy(nn,nt),xx(nn,nt,nr)
call check(b,xx)
f=dfloat(nn)
ftot=dfloat(nob)
fnt=dfloat(nt)
n1=nr+1
n2=nr+2
s2=b(n1)
g=b(n2)
u=0.0
e=0.0
if (nmueq.1) then
n3=nr+3
u=b(n3)
if (netaeq.1) then
n4=nr+4
e=b(n4)
endif
else
if (netaeq.1) then
n4=nr+3
e=b(n4)
endif
endif
sc=1.
if (ipceq.2) sc=-1.
z=u/(s2*g)**0.5
do 106 j=1,n
gx(j)=0.0
106 continue
gx(n1)=0.5*ftot/s2-0.5*f*(den(-z)/(dis(z))+z)*z/s2
gx(n2)=-.5*(ftot-f)/(1.-g)-.5*f*(den(-z)/(1.-dis(-z))+z)*z/g

do 105 i=1,nn
epr=0.0
epe=0.0
do 103 l=1,nt
if (xx(i,l,1).ne.0.0) then
ee=yy(i,l)
do 102 j=1,nb
ee=ee-b(j)*xx(i,l,j)
102 continue
epr=epr+ee*dexp(-e*(dfloat(l)-fnt))
epe=epe+dexp(-2.0*e*(dfloat(l)-fnt))
end if
103 continue
zi=(u*(1.0-g)-sc*g*epr)/(g*(1.0-g)*s2*(1.0+(epe-1.0)*g))**0.5

do 132 j=1,nb
do 134 l=1,nt
if (xx(i,l,1).ne.0.0) then
ee=yy(i,l)
do 135 k=1,nb
ee=ee-xx(i,l,k)*b(k)
135 continue

```

```

    gx(j)=gx(j)-xx(i,l,j)*ee/(s2*(1.-g))
    endif
134 continue
    xpe=0.0
    do 146 l=1,nt
    if(xx(i,l,1).ne.0.0)xpe=xpe+xx(i,l,j)*dexp(-e*(dfloat(l)-fnt))
146 continue
    d=(den(-zi)/(dis(zi))+zi)*g*xpe*sc
    gx(j)=gx(j)-d/(g*(1.0-g)*s2*(1.0+(epe-1.0)*g))**0.5
132 continue

    gx(n1)=gx(n1)+.5*(den(-zi)/(dis(zi))+zi)*zi/s2
    ss=0.0
    do 138 l=1,nt
    ee=yy(i,l)
    do 139 j=1,nb
    ee=ee-xx(i,l,j)*b(j)
139 continue
    ss=ss+ee**2
138 continue
    gx(n1)=gx(n1)-0.5*ss/((1.0-g)*s2**2)

    gx(n2)=gx(n2)+0.5*ss/((1.0-g)**2*s2)
    gx(n2)=gx(n2)+0.5*(epe-1.0)/(1.0+(epe-1.0)*g)
    d=g*(1.-g)*(1.0+(epe-1.0)*g)
    dzi=-(u+sc*epr)*d
    c=0.5*(u*(1.0-g)-sc*g*epr)
    dzi=dzi-c*((1.0-2.0*g)+(epe-1.0)*g*(2.0-3.0*g))
    dzi=dzi/(d**1.5*s2**0.5)
    gx(n2)=gx(n2)-(den(-zi)/(dis(zi))+zi)*dzi

    if (nmueq.1) then
    gx(n3)=gx(n3)+1./(s2*g)**0.5*(den(-z)/(dis(z))+z)
    d=(den(-zi)/(dis(zi))+zi)*(1.-g)
    gx(n3)=gx(n3)-d/(g*(1.-g)*s2*(1.+(epe-1.0)*g))**0.5
    end if

    if (netaeq.1) then
    de=0.0
    d=0.0
    do 152 l=1,nt
    if (xx(i,l,1).eq.1) then
    t=dfloat(l)
    de=de-2.0*(t-fnt)*dexp(-2.0*e*(t-fnt))
    ee=yy(i,l)
    do 153 j=1,nb
    ee=ee-xx(i,l,j)*b(j)
153 continue
    d=d+(t-fnt)*dexp(-e*(t-fnt))*ee
    end if
152 continue
    dd=(g*(1.0-g)*s2*(1.0+(epe-1.0)*g))
    d=d*g*dd*sc
    c=u*(1.0-g)-sc*g*epr
    c=c*0.5*g**2*(1.0-g)*s2*de
    dzi=(d-c)/dd**1.5
    gx(n4)=gx(n4)-(den(-zi)/(dis(zi))+zi)*dzi
    gx(n4)=gx(n4)+g/2.0*de/(1.0+(epe-1.0)*g)

```

```

    end if
105 continue

    ndrv=ndrv+1
    return
end

subroutine fun2(b,a,yy,xx)
c calculates the negative of the log-likelihood function of the
c TE effects model.
implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
dimension b(n),yy(nn,nt),xx(nn,nt,nr)
data pi/3.1415926/
call check(b,xx)
s2=b(nr+1)
g=b(nr+2)
ss=(g*(1.-g)*s2)**0.5
sc=1.0
if (ipc.eq.2) sc=-1.0
a=0.
do 10 i=1,nn
do 10 l=1,nt
if (xx(i,l,1).ne.0.0) then
xb=0.
do 11 j=1,nb
xb=xb+xx(i,l,j)*b(j)
11 continue
ee=(yy(i,l)-xb)
zd=0.
if (nz.ne.0) then
do 12 j=nb+1,nr
zd=zd+xx(i,l,j)*b(j)
12 continue
endif
us=(1.-g)*zd-sc*g*ee
d=zd/(g*s2)**0.5
ds=us/ss
a=a-0.5*dlog(2.*pi)-0.5*dlog(s2)-(dlog(dis(d))-dlog(dis(ds)))
+ -0.5*(ee+sc*zd)**2/s2
endif
10 continue
a=-a
nfunct=nfunct+1
return
end

subroutine der2(b,gx,yy,xx)
c calculates the first-order partial derivatives of the negative
c of the log-likelihood function of the TE effects model.
implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
dimension b(n),gx(n),yy(nn,nt),xx(nn,nt,nr)
call check(b,xx)
s2=b(nr+1)

```

```

      g=b(nr+2)
      ss=(g*(1.-g)*s2)**0.5
      sc=1.0
      if (ipc.eq.2) sc=-1.0
      do 9 j=1,n
        gx(j)=0.
9      continue
        do 10 i=1,nn
          do 10 l=1,nt
            if (xx(i,l,1).ne.0.0) then
              xb=0.
              do 11 j=1,nb
                xb=xb+xx(i,l,j)*b(j)
11      continue
                ee=(yy(i,l)-xb)
                zd=0.
                if (nz.ne.0) then
                  do 12 j=nb+1,nr
                    zd=zd+xx(i,l,j)*b(j)
12      continue
                  endif
                  us=(1.-g)*zd-sc*g*ee
                  d=zd/(g*s2)**0.5
                  ds=us/ss
                  do 13 j=1,nb
                    gx(j)=gx(j)+xx(i,l,j)*((ee+sc*zd)/s2+sc*den(ds)/dis(ds)*g/ss)
13      continue
                    if (nz.ne.0) then
                      do 14 j=nb+1,nr
                        gx(j)=gx(j)-xx(i,l,j)*((sc*ee+zd)/s2+den(d)/dis(d)/(g*s2)
+ **0.5-den(ds)/dis(ds)*(1.-g)/ss)
14      continue
                        endif
                        gx(nr+1)=gx(nr+1)-0.5/s2*(1.-(den(d)/dis(d)*d-den(ds)
+ /dis(ds)*ds)-(ee+sc*zd)**2/s2)
                        gx(nr+2)=gx(nr+2)+0.5*(den(d)/dis(d)*d/g-den(ds)/dis(ds)
+ /ss*(zd/g+sc*ee/(1.-g)))
c      gx(nr+2)=gx(nr+2)+0.5*(den(d)/dis(d)*d/g-den(ds)/dis(ds)*
c + (2.*(ee+zd)/ss+ds*(1.-2.*g)/(g*(1.-g))))
                        endif
10      continue
          do 15 j=1,n
            gx(j)=-gx(j)
15      continue
        ndrv=ndrv+1
        return
      end

```

```

double precision function den(a)
c  evaluates the n(0,1) density function.
  implicit double precision (a-h,o-z)
  common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
  data rrt2pi/ 0.3989422804/
  den=rrt2pi*dexp(-0.5*a**2)
  if (den.lt.1.0/bignum) den=1.0/bignum
  return
end

```

```

double precision function dis(x)
c  evaluates the n(0,1) distribution function.
implicit double precision (a-h,o-z)
common/five/tol,tol2,bignum,step1,gridno,igrd2,ite
dimension a(5),connor(17)
data connor
+ /8.0327350124d-17, 1.4483264644d-15, 2.4668270103d-14,
+ 3.9554295164d-13, 5.9477940136d-12, 8.3507027951d-11,
+ 1.0892221037d-9, 1.3122532964d-8, 1.4503852223d-7,
+ 1.4589169001d-6, 1.3227513228d-5, 1.0683760684d-4,
+ 7.5757575758d-4, 4.6296296296d-3, 2.3809523810d-2,
+ 0.1, 0.3333333333 /
data rrt2pi/0.3989422804/
s=x
y=s*s
if(s) 10,11,12
11 p=0.5
goto 31
10 s=-s
12 z=rrt2pi*dexp(-0.5*y)
if(s-2.5) 13,14,14
13 y=-0.5*y
p=connor(1)
do 15 l=2,17
15 p=p*y+connor(l)
p=(p*y+1.0)*x*rrt2pi+0.5
goto 31
14 a(2)=1.0
a(5)=1.0
a(3)=1.0
y=1.0/y
a(4)=1.0+y
r=2.0
19 do 17 l=1,3,2
do 18 j=1,2
k=l+j
ka=7-k
18 a(k)=a(ka)+a(k)*r*y
17 r=r+1.0
if(dabs(a(2)/a(3)-a(5)/a(4)).gt.1.0/bignum) goto 19
20 p=(a(5)/a(4))*z/x
if(x) 21,11,22
21 p=-p
goto 31
22 p=1.0-p
31 continue
if(p.lt.1.0/bignum) p=1.0/bignum
if(p.gt.(1.0-1.0/bignum)) p=1.0-1.0/bignum
dis=p
return
end

subroutine info
c  accepts instructions from the terminal or from a file and
c  also reads data from a file.
implicit double precision (a-h,o-z)
character*12 koutf,kdatf,kinf
common/eight/koutf,kdatf,kinf

```

```

common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
character chop,chst,chmu,cheta,chl
dimension yy(:,:),xx(:,:),mm(:,),sv(:,),xxd(:)
allocatable :: yy,xx,mm,sv,xxd
nmu=0
neta=0
igrid=1
nz=0
kinf='terminal'
write(6,*)
write(6,*)
write(6,*) '*****'
write(6,*) 'FRONTIER - Version 4.1'
write(6,*) '*****'
write(6,*)
write(6,*) 'by'
write(6,*) 'Tim Coelli'
write(6,*) 'Department of Econometrics'
write(6,*) 'University of New England'
write(6,*) 'Armidale, NSW, 2351'
write(6,*) 'Australia.'
write(6,*) '(email: tcoelli@metz.une.edu.au)'
write(6,*)
write(6,*) '[This software is not a commercial product.'
write(6,*) 'If you find it useful, a contribution of'
write(6,*) '$200 (US$150) to help cover some of the costs'
write(6,*) 'associated with development would be'
write(6,*) 'appreciated - but is in no way obligatory.]'
write(6,*)
write(6,*)
write(6,*) 'Do you wish to type instructions at the terminal (t)'
write(6,*) 'or use an instruction file (f) ? '
read(5,61) chop
61 format(a)
if ((chop.eq.'t').or.(chop.eq.'T')) then
write(6,*) 'enter 1 if you wish to estimate the error '
write(6,*) 'components model, or 2 for the TE effects model : '
read(5,*) im
if ((im.ne.1).and.(im.ne.2)) then
write(6,*) 'incorrect option, must be 1 or 2 - bye!'
stop
endif
write(6,*) 'enter the name of your data file : '
read(5,60) kdatf
write(6,*) 'enter a name for an output file : '
read(5,60) koutf
write(6,*) 'are you estimating a production or cost function?'
write(6,*) 'enter a 1 for production or a 2 for cost : '
read(5,*) ipc
write(6,*) 'is the dependent variable logged? (y or n) '
read(5,61) chl
if ((chl.eq.'y').or.(chl.eq.'Y')) il=1
write(6,*) 'how many cross-sections in the data ? '
read(5,*) nn
write(6,*) 'how many time-periods in the data ? '
read(5,*) nt
write(6,*) 'how many observations in total in the data ? '

```



```

read(5,*) nob
if ((nn*nt).lt.nob) then
write(6,*) ' the above number is larger than the product of the'
write(6,*) ' previous two answers - program abort!'
stop
end if
write(6,*) ' how many regressors (Xs) are there ? '
read(5,*) nb
if (im.eq.1) then
write(6,*) ' does the model include mu ? (y or n) '
read(5,61) chmu
if ((chmu.eq.'y').or.(chmu.eq.'Y')) nmu=1
if (nt.gt.1) then
write(6,*) ' does the model include eta ? (y or n) '
read(5,61) cheta
if ((cheta.eq.'y').or.(cheta.eq.'Y')) neta=1
endif
nb=nb+1
nr=nb
n=nr+2+nmu+neteta
else
write(6,*) ' does the model include delta0 ? (y or n) '
read(5,61) chmu
if ((chmu.eq.'y').or.(chmu.eq.'Y')) nmu=1
write(6,*) ' how many eff.-term regressors (Zs) are there ? '
read(5,*) nz
nz=nz+nmu
nb=nb+1
nr=nb+nz
n=nr+2
endif
allocate(sv(n))
write(6,*) ' if you do not wish the computer to select'
write(6,*) ' starting values using a grid search you must'
write(6,*) ' supply them now'
write(6,*) ' do you wish to supply starting values ? (y or n) '
read(5,61) chst
if ((chst.eq.'y').or.(chst.eq.'Y')) then
igrd=0
62 format(' enter starting value for b',i2,' : ')
do 138 i=1,nb
write(6,62) i-1
read(5,*) sv(i)
138 continue
write(6,*) ' enter starting value for sigma squared : '
read(5,*) sv(nr+1)
write(6,*) ' enter starting value for gamma : '
read(5,*) sv(nr+2)
if ((im.eq.2).and.(nz.gt.0)) then
63 format(' enter starting value for delta',i2,' : ')
do 139 i=1,nz
write(6,63) i-nmu
read(5,*) sv(nb+i)
139 continue
endif
if (im.eq.1) then
if (nmu.eq.1) then
write(6,*) ' enter starting value for mu : '

```

```

read(5,*) sv(nb+3)
if (neta.eq.1) then
write(6,*) ' enter starting value for eta : '
read(5,*) sv(nb+4)
end if
else
if (neta.eq.1) then
write(6,*) ' enter starting value for eta : '
read(5,*) sv(nb+3)
end if
end if
endif
end if
else if ((chop.eq.'f').or.(chop.eq.'F')) then
write(6,*) ' enter instruction file name : '
read(5,60) kinf
60 format(a12)
open(unit=50,file=kinf,status='old')
read(50,*) im
read(50,60) kdatf
read(50,60) koutf
read(50,*) ipc
read(50,61) chl
if ((chl.eq.'y').or.(chl.eq.'Y')) il=1
read(50,*) nn
read(50,*) nt
read(50,*) nob
if ((nn*nt).lt.nob) then
write(6,*) ' the total number of obsns exceeds the product of
write(6,*) ' the number of firms by the number of years - bye!'
stop
end if
read(50,*) nb
if (im.eq.1) then
read(50,61) chmu
read(50,61) cheta
if ((chmu.eq.'y').or.(chmu.eq.'Y')) nmu=1
if ((cheta.eq.'y').or.(cheta.eq.'Y')) neta=1
nb=nb+1
nr=nb
n=nr+2+nmu+neta
else
read(50,61) chmu
if ((chmu.eq.'y').or.(chmu.eq.'Y')) nmu=1
read(50,*) nz
nz=nz+nmu
nb=nb+1
nr=nb+nz
n=nr+2
endif
allocate (sv(n))
read(50,61) chst
if ((chst.eq.'y').or.(chst.eq.'Y')) then
igrd=0
if (im.eq.1) then
do 148 i=1,n
read(50,*) sv(i)
148 continue

```

```

else
do 152 i=1,nb
read(50,*) sv(i)
152 continue
read(50,*) sv(n-1)
read(50,*) sv(n)
if (nz.gt.0) then
do 153 i=1,nz
read(50,*) sv(nb+i)
153 continue
endif
endif
end if
else
write(6,*) ' incorrect option - must be t or f - bye!'
stop
endif
allocate(yy(nn,nt),xx(nn,nt,nr),mm(nn),xxd(nr))
open(unit=40,file=kdatf,status='old')
do 135 i=1,nn
mm(i)=0
do 135 l=1,nt
xx(i,l,1)=0.0
135 continue
do 134 k=1,nob
ndat=nr
if(im.eq.2) ndat=nr-nmu
read(40,*) fii,ftt,yyd,(xxd(j),j=2,ndat)
i=int(fii)
l=int(ftt)
mm(i)=mm(i)+1
xx(i,l,1)=1.0
yy(i,l)=yyd
do 136 j=2,nb
xx(i,l,j)=xxd(j)
136 continue
if ((im.eq.2).and.(nz.gt.0)) then
if (nmu.eq.1) xx(i,l,nb+1)=1.0
if ((nz-nmu).gt.0) then
do 154 j=nb+nmu+1,nr
xx(i,l,j)=xxd(j-nmu)
154 continue
endif
endif
if (i.lt.1) then
write(6,*) ' error - a firm number is < 1'
stop
else if (i.gt.nn) then
write(6,*) ' error - a firm number is > number of firms'
stop
else if (l.lt.1) then
write(6,*) ' error - a period number is < 1'
stop
else if (l.gt.nt) then
write(6,*) ' error - a period number is > number of periods'
stop
end if
134 continue

```

```

do 149 i=1,nn
  if (mm(i).eq.0) then
    write(6,66) i
66  format(' error - there are no observations on firm ',i6)
    stop
  end if
149 continue
  call mini(yy,xx,mm,sv)
  deallocate(yy,xx,mm,sv,xxd)
  return
end

subroutine result(yy,xx,mm,h,y,sv,ob,obse,gb,ncall)
c  presents estimates, covariance matrix, standard errors and t-ratios,
c  as well as presenting many results including estimates of technical
c  efficiency.
  implicit double precision (a-h,o-z)
  character*12 koutf,kdatf,kinf
  common/eight/koutf,kdatf,kinf
  common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
  common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
  common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
  dimension yy(nn,nt),xx(nn,nt,nr),mm(nn)
  dimension h(n,n),y(n),sv(n),ob(n),obse(n),gb(n)
  dimension mt(:)
  allocatable :: mt
  data pi/3.1415926/
  allocate(mt(nt))
  open (unit=70,file=koutf,status='unknown')
  n1=nr+1
  n2=nr+2
  n3=nr+3
  n4=nr+4
  if((nmu.eq.0).and.(neta.eq.1)) n4=nr+3
  if (ncall.eq.1) then
    write (70,401)
401  format(/,'Output from the program FRONTIER (Version 4.1)',/)
    write (70,601) kinf,kdatf
601  format('instruction file = ',a12,/, 'data file = ',7x,a12,/)
    if (im.eq.1) then
      write(70,*) 'Error Components Frontier (see B&C 1992)'
    else
      write(70,*) 'Tech. Eff. Effects Frontier (see B&C 1993)'
    endif
    if (ipc.eq.1) then
      write(70,*) 'The model is a production function'
    else
      write(70,*) 'The model is a cost function'
    endif
    if (il.eq.1) then
      write(70,*) 'The dependent variable is logged'
    else
      write(70,*) 'The dependent variable is not logged'
    endif
    fnob=dfloat(nob)
    fnb=dfloat(nb)
    os2=ob(nb+1)*(fnob-fnb)/fnob

```

```

        fxols=fnob/2.0*(dlog(2.0*pi)+dlog(os2)+1.0)
        if(igrid.eq.1) then
            write (70,403)
403    format(/'the ols estimates are :',/)
            write (70,404)
404    format(17x,'coefficient    standard-error    t-ratio',/)
            do 132 i=1,nb
                write (70,302) i-1,ob(i),obse(i),ob(i)/obse(i)
132    continue
                write (70,203) ob(nb+1)
                write(70,501) -fxols
            endif
            if(igrid.eq.0) then
                write (70,402)
402    format(/'the starting values supplied were :',/)
202    format('  beta',i2,7x,e16.8)
        82    format('  delta',i2,6x,e16.8)
203    format('  sigma-squared',e16.8)
204    format('  gamma',8x,e16.8)
205    format('  mu          ',e16.8)
206    format('  eta          ',e16.8)
            do 131 i=1,nb
                write (70,202) i-1,sv(i)
131    continue
                write (70,203) sv(n1)
                write (70,204) sv(n2)
                if ((nz.ne.0).and.(im.eq.2)) then
                    do 83 i=1,nz
                        write(70,82) i-nmu,sv(nb+i)
83    continue
                    endif
                    if (im.eq.1) then
                        if (nmueq.1) then
                            write (70,205) sv(n3)
                        else
                            write (70,*) '  mu is restricted to be zero'
                        end if
                    if (netaeq.1) then
                        write (70,206) sv(n4)
                    else
                        write (70,*) '  eta is restricted to be zero'
                    end if
                endif
                end if
                if(igrid.eq.1) then
                    write (70,405)
405    format(/,'the estimates after the grid search were :',/)
                    do 133 i=1,nb
                        write (70,202) i-1,gb(i)
133    continue
                        write (70,203) gb(n1)
                        write (70,204) gb(n2)
                        if ((nz.ne.0).and.(im.eq.2)) then
                            do 84 i=1,nz
                                write(70,82) i-nmu,gb(nb+i)
84    continue
                            endif
                            if (im.eq.1) then

```

```

    if (nmu.eq.1) then
    write (70,205) gb(n3)
    else
    write (70,*) ' mu is restricted to be zero'
    end if
    if (neta.eq.1) then
    write (70,206) gb(n4)
    else
    write (70,*) ' eta is restricted to be zero'
    end if
    endif
    end if

    else

    write(6,151) koutf
151  format(' sending output to: ',a12)
    write (70,406)
406  format('//, 'the final mle estimates are : ',/)
    write (70,404)
302  format(' beta',i2,7x,3e16.8)
    81  format(' delta',i2,6x,3e16.8)
303  format(' sigma-squared',3e16.8)
304  format(' gamma',8x,3e16.8)
305  format(' mu      ',3e16.8)
306  format(' eta      ',3e16.8)
    do 134 i=1,nb
    write (70,302) i-1,y(i),h(i,i)**0.5,y(i)/h(i,i)**0.5
134  continue
    write (70,303) y(n1),h(n1,n1)**0.5,y(n1)/h(n1,n1)**0.5
    write (70,304) y(n2),h(n2,n2)**0.5,y(n2)/h(n2,n2)**0.5
    if ((nz.ne.0).and.(im.eq.2)) then
    do 199 i=nb+1,nb+nz
    write (70,81) i-nmu-nb,y(i),h(i,i)**0.5,y(i)/h(i,i)**0.5
199  continue
    endif
    if (im.eq.1) then
    if (nmu.eq.1) then
    write (70,305) y(n3),h(n3,n3)**0.5,y(n3)/h(n3,n3)**0.5
    else
    write (70,*) ' mu is restricted to be zero'
    end if
    if (neta.eq.1) then
    write (70,306) y(n4),h(n4,n4)**0.5,y(n4)/h(n4,n4)**0.5
    else
    write (70,*) ' eta is restricted to be zero'
    end if
    endif
    write (70,501) -fx
501  format(/, 'log likelihood function = ',e16.8)
    if((fx-fxols).gt.0) then
    write(70,422)
422  format(/, 'the likelihood value is less than that obtained',
    + /, 'using ols! - try again using different starting values')
    else
    chi=2.0*dabs(fx-fxols)
    write(70,5011) chi
5011 format(/, 'LR test of the one-sided error = ',e16.8)

```

```

        if (im.eq.1) idf=nmu+neta+1
        if (im.eq.2) idf=nz+1
        write(70,5012) idf
5012 format('with number of restrictions = ',i1)
        write(70,*) '[note that this statistic has a mixed chi-square'
+ , ' distribution]'
        end if
        write (70,502) iter
502 format(/,'number of iterations = ', i6)
        write(70,420) maxit
420 format(/,'(maximum number of iterations set at : ',i6,')')
        write(70,513) nn
513 format(/,'number of cross-sections = ',i6)
        write(70,514) nt
514 format(/,'number of time periods = ',i6)
        write(70,515) nob
515 format(/,'total number of observations = ',i6)
        write(70,516) nn*nt-nob
516 format(/,'thus there are: ',i6,' obsns not in the panel')

        write (70,58)
58 format(//,'covariance matrix :',/)
52 format(5e16.8)
        do 135 i=1,n
        nc=1
314 write(70,52) (h(i,j),j=nc,min(n,nc+4))
        nc=nc+5
        if(nc.le.n) goto 314
135 continue

        if((ite.eq.1).and.(ipc.eq.1)) write(70,503)
503 format(///,'technical efficiency estimates :',/)
        if((ite.eq.1).and.(ipc.eq.2)) write(70,504)
504 format(///,'cost efficiency estimates :',/)
67 format(//,1x,'mean efficiency = ',e16.8,///)
68 format(/,' firm year eff.-est.',/)
69 format(2x,2i6,9x,e16.8)
167 format(//,1x,'mean eff. in year' ,i4,' =',e16.8,/)
168 format(/,' firm eff.-est.',/)
169 format(2x,i6,9x,e16.8)
        sc=1.
        if(ipc.eq.2) sc=-1.

        if (im.eq.1) then
        s2=y(nb+1)
        g=y(nb+2)
        u=0.0
        e=0.0
        if (nmueq.1) then
        u=y(nb+3)
        if (neta.eq.1) e=y(nb+4)
        else
        if (neta.eq.1) e=y(nb+3)
        endif
524 format(//,'efficiency estimates for year ',i6,' :')
505 format(2x,i6,4x,'no observation in this period')
        fnt=dfloat(nt)
        xbb=0.

```

```

te=0.
ncount=0
ntt=nt
if (neta.eq.0) ntt=1
do 138 l=1,ntt
t=dfloat(l)
eta=dexp(-e*(t-fnt))
if (ite.eq.1) then
if (neta.eq.1) write(70,524) l
write(70,168)
endif
do 136 i=1,nn
if ((xx(i,1,1).ne.0.0).or.(neta.eq.0)) then
epr=0.0
epe=0.0
do 103 k=1,nt
if (xx(i,k,1).ne.0.0) then
ee=yy(i,k)
do 102 j=1,nb
ee=ee-y(j)*xx(i,k,j)
102 continue
xb=yy(i,k)-ee
xbb=xbb+xb
epr=epr+ee*dexp(-e*(dfloat(k)-fnt))
epe=epe+dexp(-2.0*e*(dfloat(k)-fnt))
end if
103 continue
fi=(u*(1.0-g)-g*epr)/(1.0+(epe-1.0)*g)
si2=g*(1.0-g)*s2/(1.0+(epe-1.0)*g)
si=si2**0.5
if (il.eq.1) then
tei=(1.0-dis(sc*si*eta-fi/si))/(dis(fi/si))
tei=tei*dexp(-fi*eta*sc+0.5*si2*eta**2)
else
tei=fi+si*den(fi/si)/dis(fi/si)
tei=1.-sc*(eta*tei/(xb/dfloat(mm(i))))
endif
if ((ipc.eq.1).and.(tei.gt.1.0)) tei=1.0
if ((ipc.eq.2).and.(tei.lt.1.0)) tei=1.0
te=te+tei
ncount=ncount+1
endif
if (ite.eq.1) then
if ((xx(i,1,1).ne.0.).or.(neta.eq.0)) then
write(70,169) i,tei
else
write(70,505) i
endif
endif
136 continue
if (neta.eq.1) write(70,167) l,te/dfloat(ncount)
if (neta.eq.0) write(70,67) te/dfloat(ncount)
te=0.
ncount=0
138 continue
else

```



```

        if(ite.eq.1) write(70,68)
        s2=y(nr+1)
        g=y(nr+2)
        te=0.
        ss=(g*(1.-g)*s2)**0.5
        do 10 l=1,nt
        do 10 i=1,nn
        if (xx(i,l,1).ne.0.0) then
        xb=0.
        do 11 j=1,nb
        xb=xb+xx(i,l,j)*y(j)
11 continue
        zd=0.
        if (nz.ne.0) then
        do 12 j=nb+1,nr
        zd=zd+xx(i,l,j)*y(j)
12 continue
        endif
        us=(1.-g)*zd-g*(yy(i,l)-xb)
        ds=us/ss
        if (il.eq.1) then
        tei=dexp(-sc*us+0.5*ss**2)*dis(ds-sc*ss)/dis(ds)
        else
        tei=1.-sc*(us+ss*den(ds)/dis(ds))/xb
        endif
        if ((ipc.eq.1).and.(tei.gt.1.0)) tei=1.0
        if ((ipc.eq.2).and.(tei.lt.1.0)) tei=1.0
        if (ite.eq.1) write(70,69) i,l,tei
        te=te+tei
        endif
10 continue
        te=te/dfloat(nob)
        write(70,67) te
        endif

        if ((nt.gt.1).and.(nt.le.100)) then
        write(70,441)
441 format(///,'summary of panel of observations: ',/,
+ '(1 = observed, 0 = not observed)',/)
        do 449 l=1,nt
        mt(l)=1
449 continue
        write(70,442) (mt(l),l=1,nt)
442 format(' t:',100i4)
        write(70,*) ' n'
443 format(102i4)
        do 450 i=1,nn
        write(70,443) i,(int(xx(i,l,1)),l=1,nt),mm(i)
450 continue
        do 451 l=1,nt
        mt(l)=0
        do 451 i=1,nn
        mt(l)=mt(l)+int(xx(i,l,1))
451 continue
        write(70,444) (mt(l),l=1,nt),nob
444 format(/,4x,101i4)
        write(70,445)
445 format(////)

```

```

endif
endif
deallocate(mt)
return
end

c      subroutine grid(x,y,yy,xx,ob,gb)
      does a grid search across gamma
      implicit double precision (a-h,o-z)
      common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
      common/three/n,nfunct,ndrv,iter,indic,iprint,igrd,maxit
      common/five/tol,tol2,bignum,step1,gridno,igrd2,ite
      dimension x(n),y(n),yy(nn,nt),xx(nn,nt,nr),ob(n),gb(n)
      data pi/3.1415926/
      n1=nr+1
      n2=nr+2
      sc=1.0
      if (ipc.eq.2) sc=-1.0
      var=ob(nb+1)*float(nob-nb)/float(nob)
      b0=ob(1)
      do 131 i=1,nb+1
      y(i)=ob(i)
131  continue
      do 132 i=nb+1,nr
      y(i)=0.
132  continue
      fx=bignum
      y6b=gridno
      y6t=1.0-gridno
      do 137 y6=y6b,y6t,gridno
      y(n2)=y6
      y(n1)=var/(1.-2.*y(n2)/pi)
      c=(y(n2)*y(n1)*2/pi)**0.5
      y(1)=b0+c*sc
      if (im.eq.1) call fun1(y,fy,yy,xx)
      if (im.eq.2) call fun2(y,fy,yy,xx)
      if(fy.lt.fx) then
      fx=fy
      do 138 i=1,n
      x(i)=y(i)
138  continue
      end if
137  continue
      if(igrd2.eq.1) then
      bb1=x(n2)-gridno/2.0
      bb2=x(n2)+gridno/2.0
      bb3=gridno/10.0
      do 140 y6=bb1,bb2,bb3
      y(n2)=y6
      y(n1)=var/(1.-2.*y(n2)/pi)
      c=(y(n2)*y(n1)*2/pi)**0.5
      y(1)=b0+c*sc
      if (im.eq.1) call fun1(y,fy,yy,xx)
      if (im.eq.2) call fun2(y,fy,yy,xx)
      if(fy.lt.fx) then
      fx=fy

```

```

        do 141 i=1,n
        x(i)=y(i)
141  continue
        end if
140  continue
        end if
        do 142 i=1,n
        gb(i)=x(i)
        y(i)=x(i)
142  continue
        fy=fx
        return
        end

```

```

c      subroutine invert(xx,n)
      finds the inverse of a given matrix.
      implicit double precision (a-h,o-z)
      common/five/tol,tol2,bignum,step1,gridno,igrid2,ite
      dimension xx(n,n)
      dimension ipiv(:)
      allocatable :: ipiv
      allocate(ipiv(n))
      do 1 i=1,n
1      ipiv(i)=0
        do 11 i=1,n
          amax=0.
          do 5 j=1,n
            if(ipiv(j))2,2,5
2          if(dabs(xx(j,j))-amax) 4,4,3
3          icol=j
            amax=dabs(xx(j,j))
4          continue
5          continue
            ipiv(icol)=1
            if(amax-1.0/bignum)6,6,7
6          write(6,*) 'singular matrix'
            stop
7          continue
            amax=xx(icol,icol)
            xx(icol,icol)=1.0
            do 8 k=1,n
8          xx(icol,k)=xx(icol,k)/amax
            do 11 j=1,n
              if(j-icol)9,11,9
9          amax=xx(j,icol)
              xx(j,icol)=0.
              do 10 k=1,n
10         xx(j,k)=xx(j,k)-xx(icol,k)*amax
11         continue
              deallocate(ipiv)
              return
            end

```

```

c      subroutine ols(ob,obse,yy,xx)
      calculates the ols estimates and their standard errors.

```

```

implicit double precision (a-h,o-z)
common/one/fx,fy,nn,nz,nb,nr,nt,nob,nmu,neta,ipc,im,il
common/three/n,nfunct,ndrv,iter,indic,iprint,igrid,maxit
dimension ob(n),obse(n),yy(nn,nt),xx(nn,nt,nr)
dimension xpx(:,,:),xpy(:,),mx(:)
allocatable :: xpx,xpy,mx
allocate(xpx(nb,nb),xpy(nb),mx(nb))
c   calculate x'x and x'y
do 131 k=1,nb
do 132 j=1,nb
xpx(k,j)=0.0
do 132 i=1,nn
do 132 l=1,nt
if (xx(i,l,1).ne.0.0) xpx(k,j)=xpx(k,j)+xx(i,l,k)*xx(i,l,j)
132 continue
xpy(k)=0.0
do 131 i=1,nn
do 131 l=1,nt
if (xx(i,l,1).ne.0.0) xpy(k)=xpy(k)+xx(i,l,k)*yy(i,l)
131 continue
c   determine correct scaling for x'x
do 120 k=1,nb
h=(1.0-dlog10(xpx(k,k)))/2.0
if (h.lt.0.0) goto 121
mx(k)=h
goto 120
121 mx(k)=h-1
120 continue
c   scale, invert and then scale back
is=0
123 is=is+1
do 122 k=1,nb
do 122 j=1,nb
xpx(k,j)=xpx(k,j)*10.0**(mx(k)+mx(j))
122 continue
if (is.eq.1) then
call invert(xpx,nb)
goto 123
endif
c   calculate b=inv(x'x)x'y
do 133 k=1,nb
ob(k)=0.0
do 133 j=1,nb
ob(k)=ob(k)+xpx(k,j)*xpy(j)
133 continue
ss=0.0
do 134 i=1,nn
do 134 l=1,nt
if (xx(i,l,1).ne.0.0) then
ee=yy(i,l)
do 135 k=1,nb
ee=ee-xx(i,l,k)*ob(k)
135 continue
ss=ss+ee**2
endif
134 continue
ob(nb+1)=ss/dfloat(nob-nb)
do 136 k=1,nb

```

Appendix 6: Asymptotic Standard Errors of the COLS Estimators

The derivation of the standard errors of the COLS estimators follows the method used in the Appendix of Olsen, Schmidt and Waldman (1980). Consider the composed error disturbance term $\epsilon = V - U$, where $V \sim N(0, \sigma_V^2)$ and $U \sim N(0, \sigma^2)$. Aigner, Lovell and Schmidt (1977) show that the expectation of ϵ is

$$E(\epsilon) = \mu_1 = -\sigma\sqrt{2/\pi} \quad (\text{A6.1})$$

and Olsen, Schmidt and Waldman (1980) present the first six central moments of ϵ in terms of σ_V^2 and σ^2 . Their results can also be expressed in terms of σ_S^2 and γ by noting that $\sigma_V^2 = (1-\gamma)\sigma_S^2$ and $\sigma^2 = \gamma\sigma_S^2$. Substitution and rearrangement of terms yields

$$\mu_1 = -\sqrt{2\gamma\sigma_S^2/\pi} \quad (\text{A6.2})$$

$$\mu_2 = \sigma_S^2 \left[(1-\gamma) + \gamma \frac{(\pi-2)}{\pi} \right] \quad (\text{A6.3})$$

$$\mu_3 = \sigma_S^3 \left[\sqrt{\frac{2}{\pi}} \left(1 - \frac{4}{\pi} \right) \gamma^{3/2} \right] \quad (\text{A6.4})$$

$$\mu_4 = \sigma_S^4 \left[3(1-\gamma)^2 + 6 \frac{(\pi-2)}{\pi} \gamma(1-\gamma) + \left(3 - \frac{4}{\pi} - \frac{12}{\pi^2} \right) \gamma^2 \right] \quad (\text{A6.5})$$

$$\mu_5 = \sigma_S^5 \gamma^{3/2} \sqrt{\frac{2}{\pi}} \left[10 \left(1 - \frac{4}{\pi} \right) (1-\gamma) + \left(7 - \frac{20}{\pi} - \frac{16}{\pi^2} \right) \gamma \right] \quad (\text{A6.6})$$

$$\begin{aligned} \mu_6 = \sigma_S^6 & \left[15(1-\gamma)^3 + 45 \frac{(\pi-2)}{\pi} (1-\gamma)^2 \gamma + 15 \left(3 - \frac{4}{\pi} - \frac{12}{\pi^2} \right) (1-\gamma) \gamma^2 \right. \\ & \left. + \left(15 - \frac{6}{\pi} - \frac{100}{\pi^2} - \frac{40}{\pi^3} \right) \gamma^3 \right] \quad (\text{A6.7}) \end{aligned}$$

These expressions will be used in the following asymptotic variances and covariances of the second and third sample central moments:

$$V(m_2) = \frac{1}{N}(\mu_4 - \mu_2^2) \quad (A6.8)$$

$$V(m_3) = \frac{1}{N}(\mu_6 - \mu_3^2 - 6\mu_2\mu_4 + 9\mu_2^3) \quad (A6.9)$$

$$\text{Cov}(m_2, m_3) = \frac{1}{N}(\mu_5 - 4\mu_2\mu_3). \quad (A6.10)$$

Before these expressions can be of use, we must re-express the COLS estimators of γ and β_0 in equations (8.4) and (8.5) (in Chapter 8) to be functions of the second and third moments. These become:

$$\hat{\gamma} = \left\{ m_2 \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-2/3} + \frac{2}{\pi} \right\}^{-1} \quad (A6.11)$$

$$\hat{\beta}_0 = \hat{\beta}_0(\text{OLS}) + \sqrt{\frac{2}{\pi}} \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{1/3} \quad (A6.12)$$

and the estimator of σ_s^2 from equation (8.3) remains:

$$\hat{\sigma}_s^2 = m_2 + \frac{2}{\pi} \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{2/3}. \quad (A6.13)$$

We now use Taylor-series expansions, truncated after the first-order terms, to obtain approximations of the asymptotic standard errors of the COLS estimators. To do this for $\hat{\beta}_0$ we need to specify the asymptotic covariance between $\hat{\beta}_0(\text{OLS})$ and the third sample moment. Olsen, Schmidt and Waldman (1980) note that it is equal to zero when the model contains an intercept and no other regressors. Assuming this result extends to the general case, we obtain:

$$V(\hat{\beta}_0) = \left[\frac{\partial \hat{\beta}_0}{\partial \hat{\beta}_0(\text{OLS})} \right]^2 \cdot V(\hat{\beta}_0(\text{OLS})) + \left[\frac{\partial \hat{\beta}_0}{\partial m_3} \right]^2 \cdot V(m_3) \quad (A6.14)$$

$$V(\hat{\sigma}_s^2) = \left[\frac{\partial \hat{\sigma}_s^2}{\partial m_2} \right]^2 \cdot V(m_2) + \left[\frac{\partial \hat{\sigma}_s^2}{\partial m_3} \right]^2 \cdot V(m_3) + \left[\frac{\partial \hat{\sigma}_s^2}{\partial m_2} \right] \left[\frac{\partial \hat{\sigma}_s^2}{\partial m_3} \right] \text{Cov}(m_2, m_3) \quad (A6.15)$$

$$V(\hat{\gamma}) = \left[\frac{\partial \hat{\gamma}}{\partial m_2} \right]^2 \cdot V(m_2) + \left[\frac{\partial \hat{\gamma}}{\partial m_3} \right]^2 \cdot V(m_3) + \left[\frac{\partial \hat{\gamma}}{\partial m_2} \right] \left[\frac{\partial \hat{\gamma}}{\partial m_3} \right] \text{Cov}(m_2, m_3), \quad (\text{A6.16})$$

where

$$\left[\frac{\partial \hat{\beta}_0}{\partial \hat{\beta}_0(\text{OLS})} \right] = 1 \quad (\text{A6.17})$$

$$\left[\frac{\partial \hat{\beta}_0}{\partial m_3} \right] = \frac{\pi}{(\pi-4)} \cdot \frac{1}{3} \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-2/3} \quad (\text{A6.18})$$

$$\left[\frac{\partial \hat{\sigma}_s^2}{\partial m_2} \right] = 1 \quad (\text{A6.19})$$

$$\left[\frac{\partial \hat{\sigma}_s^2}{\partial m_3} \right] = \sqrt{\frac{2}{\pi}} \cdot \frac{\pi}{(\pi-4)} \cdot \frac{2}{3} \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-1/3} \quad (\text{A6.20})$$

$$\left[\frac{\partial \hat{\gamma}}{\partial m_2} \right] = - \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-2/3} \left\{ m_3 \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-2/3} + \frac{2}{\pi} \right\}^{-2} \quad (\text{A6.21})$$

$$\left[\frac{\partial \hat{\gamma}}{\partial m_3} \right] = m_2 \sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} \cdot \frac{2}{3} \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-5/3} \left\{ m_3 \left[\sqrt{\frac{\pi}{2}} \cdot \frac{\pi}{(\pi-4)} m_3 \right]^{-2/3} + \frac{2}{\pi} \right\}^{-2} \quad (\text{A6.22})$$

Appendix 7: SHAZAM Code For COLS

This code assumes a data file (eg.dta) which contains three columns of data, each of length 60, on output and two inputs. The code obtains OLS estimates of a Cobb-Douglas production function. It then calculates the third-moment test statistic before calculating the COLS estimates and their standard errors.

```
sample 1 60
read(eg.dta) y x1 x2
genr ly=log(y)
genr lx1=log(x1)
genr lx2=log(x2)
ols ly lx1 lx2/resid=e coef=b stderr=se
gen1 b0ols=b:3
gen1 vb0ols=(se:3)**2
genr e2=e**2
stat e2/mean=m2
genr e3=e**3
stat e3/mean=m3
gen1 n=60
gen1 m3t=m3/sqrt(6*m2**3/n)
* third moment test:
print m3t
gen1 pi=3.142
gen1 s2=m2+2/pi*(sqrt(pi/2)*pi/(pi-4)*m3)**(2/3)
gen1 g=(sqrt(pi/2)*pi/(pi-4)*m3)**(2/3)/s2
gen1 b0=b0ols+sqrt(2*g*s2/pi)
* COLS estimates:
print s2 g b0
gen1 u2=s2*((1-g)+g*(pi-2)/pi)
gen1 u3=s2**(3/2)*(sqrt(2/pi)*(1-4/pi)*g**(3/2))
gen1 u4=s2**2*(3*(1-g)**2+6*(pi-2)/pi*g*(1-g)+(3-4/pi-12/pi**2)*g**2)
gen1 u5=s2**(5/2)*g**(3/2)*sqrt(2/pi)*(10*(1-4/pi)*(1-g)+(7-20/pi-16/pi**2)*g)
gen1 u6=s2**3*(15*(1-g)**3+45*(pi-2)/pi*(1-g)**2*g &
+15*(3-4/pi-12/pi**2)*(1-g)*g**2+(15-6/pi-100/pi**2-40/pi**3)*g**3)
gen1 vm2=1/n*(u4-u2**2)
gen1 vm3=1/n*(u6-u3**2-6*u2*u4+9*u2**3)
gen1 cov=1/n*(u5-4*u2*u3)
gen1 tmp=sqrt(pi/2)*pi/(pi-4)*m3
gen1 db0dm3=pi/(pi-4)/3*tmp**(-2/3)
gen1 ds2dm3=sqrt(2/pi)*pi/(pi-4)*2/3*tmp**(-1/3)
gen1 dgdm2=-(tmp**(-2/3))*(tmp**(-2/3)+2/pi)**(-2)
gen1 dgdm3=m2*sqrt(pi/2)*pi/(pi-4)*2/3*tmp**(-5/3)*(m2*tmp**(-2/3)+2/pi)**(-2)
gen1 seb0=sqrt(vb0ols+db0dm3**2*vm3)
gen1 ses2=sqrt(vm2+ds2dm3**2*vm3+2*ds2dm3*cov)
gen1 seg=sqrt(dgdm2**2*vm2+dgdm3**2*vm3+2*dgdm2*dgdm3*cov)
* standard errors of COLS estimates:
print seb0 ses2 seg
stop
```


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